## SYSTEMTEORI - THE KALMAN FILTER

## 1. Examples

1.1. **7.5.** Let y be a stochastic process given by the system:

$$\begin{aligned} x(t+1) &= Ax(t) + Bv(t) \\ y(t) &= Cx(t) + Dv(t) \end{aligned}$$

where x(0) and v(t) satisfies the usual assumptions. Determine a Kalman filter for this case.

Note that in the normal Kalman filter framework we have the following assumptions:

1:  $x(0) = x_0$ 2:  $Ex_0 x_0^T = P_0$ 3:  $Ex_0 = 0$ 4:  $Ev(t)v^T(s) = \delta_{t,s}$ 5:  $x_0$  and v(t) uncorrelated gaussian processes.

We will determine a recursive algorithm to compute the prediction  $\hat{x}(t+1)$  given  $\hat{x}(t)$  and y(t). We will follow the same method as in chapter 9 of the compendium, and therefore we will use the same notation. We have that:

$$\begin{aligned} \hat{x}(t+1) &= E^{H_t(y)}x(t+1) = \{\text{remember that } H_t(y) = H_{t-1} \oplus [\tilde{y}(t)] \} \\ &= E^{H_{t-1}(y)}x(t+1) + E^{[\tilde{y}(t)]}x(t+1) \\ &= E^{H_{t-1}(y)}(Ax(t) + Bv(t)) + E^{\tilde{y}(t)}x(t+1) = A\hat{x}(t) + K(t)\tilde{y}(t) \end{aligned}$$

First, let us write  $\tilde{y}(t)$  in a different way:

$$\begin{split} \tilde{y}(t) &= y(t) - E^{H_{t-1}(y)}y(t) = y(t) - E^{H_{t-1}y}(Cx(t) + Dv(t)) \\ &= y(t) - C\hat{x}(t) = C\tilde{x}(t) + Dv(t) \end{split}$$

Then, we can determine K(t) with the Projection Theorem. Note that since  $E^{[\tilde{y}(t)]}x(t+1) = K(t)\tilde{y}(t)$ , we have by the Projection Theorem that

$$x(t+1) - K(t)\tilde{y}(t) \perp \tilde{y}(t),$$

and hence

$$\begin{split} E\left[(x(t+1) - K(t)\tilde{y}(t))\tilde{y}(t)^{T}\right] &= 0\\ \Rightarrow \quad Ex(t+1)\tilde{y}(t) = K(t)E\tilde{y}(t)\tilde{y}(t)^{T}\\ \Rightarrow \quad K(t) = Ex(t+1)\tilde{y}(t)^{T}[E\tilde{y}(t)\tilde{y}(t)^{T}]^{-1} \end{split}$$

We need to determine the quantities  $Ex(t+1)\tilde{y}(t)^T$  and  $E\tilde{y}(t)\tilde{y}(t)^T$ :

$$\begin{aligned} Ex(t+1)\tilde{y}(t)^T &= E(Ax(t) + Bv(t))(\tilde{x}(t)^T C^T + v(t)^T D^T) \\ &= E(A\tilde{x}(t) + A\hat{x}(t) + Bv(t))(\tilde{x}(t)^T C^T + v(t)^T D^T) \\ &= AE\tilde{x}(t)\tilde{x}(t)^T C^T + BEv(t)v(t)^T D^T \\ &= AP(t)C^T + BD^T \end{aligned}$$

where  $P(t) \triangleq E\tilde{x}(t)\tilde{x}(t)^T$ . Then, we have that:

$$\begin{split} E\tilde{y}(t)\tilde{y}(t)^T &= E(C\tilde{x}(t) + Dv(t))(\tilde{x}(t)^T C^T + v(t)^T D^T) \\ &= CE\tilde{x}(t)\tilde{x}(t)^T C^T + DEv(t)v(t)^T D^T \\ &= CP(t)C^T + DD^T \end{split}$$

So, we have the following:

$$K(t) = (AP(t)C^{T} + BD^{T})(CP(t)C^{T} + DD^{T})^{-1}$$

Finally, we need to determine a recursive equation for P(t). First, let us determine the dynamic of  $\tilde{x}(t)$ . We have that:

$$\begin{aligned} \tilde{x}(t+1) &= x(t+1) - \hat{x}(t+1) \\ &= Ax(t) + Bv(t) - A\hat{x}(t) - K(t)(C\tilde{x}(t) + Dv(t)) \\ &= (A - K(t)C)\tilde{x}(t) + (B - K(t)D)v(t) \end{aligned}$$

Hence, the dynamic for P(t+1) is:

$$P(t+1) = E\tilde{x}(t+1)\tilde{x}^{T}(t+1)$$
  
=  $E[(A - K(t)C)\tilde{x}(t) + (B - K(t)D)v(t)][\tilde{x}(t)^{T}(A - K(t)C)^{T} + v(t)^{T}(B - K(t)D)^{T}]$   
=  $(A - K(t)C)P(t)(A - K(t)C)^{T} + (B - K(t)D)(B - K(t)D)^{T},$ 

which after some manipulations and by replacing K(t) turns out to be  $P(t+1) = AP(t)A^T - (AP(t)C^T + BD^T)(CP(t)C^T + DD^T)^{-1}(AP(t)C^T + BD^T)^T + BB^T.$ 

1.2. 7.1. Consider the following system:

$$\begin{aligned} x(t+1) &= Ax(t) + Bv(t) \\ y(t) &= Cx(t) \end{aligned}$$

with

$$\begin{aligned} x(0) &= x_0\\ Ev(t)v^T(s) &= \delta_{t,s}I\\ Ex_0x_0^T &= P_0 \end{aligned}$$

and  $x_0$  and v(t) uncorrelated, zero mean Gaussian variables.

- a: Determine the Kalman filter for this case, and show that the usual Kalman
  - filter converges at this one when  $D \to 0$ .
  - **b:** Use the previous result for the following scalar case:

$$x(t+1) = \frac{1}{2}x(t) + v(t)$$
$$y(t) = x(t)$$

**a:** We will obtain the Kalman filter in the same way as in the previous exercise. We have that:

$$\hat{x}(t+1) = E^{H_t(y)}x(t+1) = E^{H_{t-1}(y)}x(t+1) + E^{[\hat{y}]}x(t+1) = A\hat{x}(t) + K_t\tilde{y}(t)$$

where:

$$\tilde{y}(t) = y(t) - E^{H_{t-1}(y)}y(t) = y(t) - E^{H_{t-1}}(Cx(t)) = y(t) - C\hat{x}(t) = C\tilde{x}(t)$$

The projection theorem gives:

$$K(t) = Ex(t+1)\tilde{y}(t)^T (E\tilde{y}(t)\tilde{y}(t)^T)^{-1}$$

Besides, we have that:

$$Ex(t+1)\tilde{y}(t)^{T} = E\tilde{x}(t+1)\tilde{y}(t)^{T} + E\hat{x}(t+1)\tilde{y}(t)^{T} = E(A\tilde{x}(t) + Bv(t))(\tilde{x}(t)^{T}C^{T}) = AP(t)C^{T}$$

Then, we have that:

$$E\tilde{y}(t)\tilde{y}(t)^T = E(C\tilde{x}(t))(\tilde{x}(t)^T C^T) = CE\tilde{x}(t)\tilde{x}(t)^T C^T = CP_t C^T$$

So, we have the following:

$$K(t) = (AP(t)C^T)(CP(t)C^T)^{-1}$$

Finally, we need to determine a recursive equation for P(t). First, let us determine the dynamic of  $\tilde{x}(t)$ . We have that:

$$\tilde{x}(t+1) = x(t+1) - \hat{x}(t+1) = Ax(t) + Bv(t) - A\hat{x}(t) - K(t)(C\tilde{x}(t)) = (A - K(t)C)\tilde{x}(t) + Bv(t)$$

Hence, the dynamic for P(t+1) is:

$$P(t+1) = E\tilde{x}(t+1)\tilde{x}^{T}(t+1) = (A - K(t)C)P(t)(A - K(t)C)^{T} + BB^{T}$$

If we insert the value for K(t) we have found previously, we get that:

$$P(t+1) = AP(t)A^{T} - AP(t)C^{T}(CP(t)C^{T})^{-1}CP(t)A^{T} + BB^{T}$$

which is exactly the usual Kalman filter when  $D \to 0$  b: In this case we have

$$A=\frac{1}{2}, \quad B=1, \quad C=1$$

By applying the previous results we get that:

$$K(t) = \frac{1}{2}P(t)P(t)^{-1} = \frac{1}{2}$$

and

$$P(t+1) = \left(\frac{1}{4}P(t) - \frac{1}{4}\right)P(t) + 1 = 1$$

So, the optimal filter is:

$$\hat{x}(t+1) = \frac{1}{2}y(t)$$

1.3. Es. 3. Consider two unknown but correlated constants,  $x_1$  and  $x_2$ . We wish to determine the improvement in the knowledge of  $x_1$  which is possible through a single noisy measurement of  $x_2$ . The vector and matrix quantities of interest are:

$$x = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}, \quad P_0 = \begin{bmatrix} \sigma_1^2 & \sigma_{12}^2 \\ \sigma_{12}^2 & \sigma_2^2 \end{bmatrix}$$

and let r be the covariance of the measurement noise.

We can rewrite the problem in the following way:

$$\begin{aligned} x(t+1) &= x(t) \\ y(t) &= \begin{bmatrix} 0 & 1 \end{bmatrix} x(t) + rv(t) \end{aligned}$$

 $P_0$  is the covariance matrix describing the uncertainty in x before the measurement. That is,  $\sigma_1^2$  is the initial mean square error in knowledge of  $x_1$ ,  $\sigma_2^2$  is the initial mean square error in knowledge of  $x_2$ , and  $\sigma_{12}$  measures the corresponding cross-correlation.

After one measurement, the updated covariance matrix, P(1), is given by:

$$P(1) = \begin{bmatrix} \sigma_1^2 \left(\frac{\sigma_2^2 (1-\rho^2) + r^2}{\sigma_2^2 + r^2}\right) & \sigma_{12}^2 \left(\frac{r^2}{\sigma_2^2 + r^2}\right) \\ \sigma_{12}^2 \left(\frac{r^2}{\sigma_2^2 + r^2}\right) & \sigma_2^2 \left(\frac{r^2}{\sigma_2^2 + r^2}\right) \end{bmatrix}$$

where  $\rho$  is the correlation coefficient, defined as:

$$\rho = \frac{\sigma_{12}^2}{\sigma_1 \sigma_2}$$

A few limiting cases are worth examining. First, in the case where the measurement is perfect, i.e. r = 0, the final uncertainty in the estimate of  $x_2$ ,  $P_{22}(1)$ , is zero. Also, when  $\rho = 0$ , the final uncertainty in the estimate of  $x_1$  is equal to the initial uncertainty: nothing can be learned from the measurement in this case. Finally, in the case where  $\rho = \pm 1$ , the final uncertainty in the estimate of  $x_1$  is given by:

$$P_{11}(1) = \sigma_1^2(\frac{1}{1 + \sigma_2^2/r^2})$$

and the amount of information gained (i.e. the reduction in  $P_{11}(1)$ ) depends on the ration of the initial mean square error in the knowledge of  $x_2$  to the mean square error in the measurement of  $x_2$ . All these results are clearly very intuitive.

1.4. Es. 4. Consider the *n*-dimensional Brownian motion w(t):

- (1) w(t) continuous a.s.
- (2) w(0) = 0
- (3) independent increments
- (4)  $w(t) w(s) \in N(0, t s).$

Suppose that we measure w(t) at the time steps: t = 1, 2, ... with a measurement noise of variance  $\sigma^2$ .

- **a:** Determine the optimal estimator of w(t) and an equation for the steady state predictor.
- **b**: Consider the one dimensional case, and let  $\sigma = 1/2$ . Plot a realization of the process, of the measurement and of the optimal estimation.
- c: Vary  $\sigma$  and plot the Kalman gain together with the corresponding steadystate Kalman gain.

d: Consider the 2-dimensional case with

$$\sigma = \left[ \begin{array}{cc} 1 & 0\\ 0 & 1/10 \end{array} \right]$$

Plot a realization of the process, the measurement and the optimal estimation.

**a:** First, we need a discrete model for the Brownian motion, and the measurement:

$$\begin{aligned} x(t+1) &= x(t) + u(t) \\ y(t) &= x(t) + \sigma v(t), \end{aligned}$$

where x(t) = w(t). Hence, we have a standard formulation of the estimation problem with:

$$A = I \quad B = I$$
$$C = I \quad D = \sigma$$

Besides, we know that the initial condition is  $x_0 = 0$  and we have that:

$$Ex_0 x_0^T = P_0 = 0$$

The solution is given by:

$$K(t) = AP(t)C^{T}(CP(t)C^{T} + DD^{T})^{-1}$$
  
=  $P(t)(P(t) + \sigma^{2}I)^{-1}$ 

$$P(t+1) = AP(t)A^{T} - AP(t)C^{T}(CP(t)C^{T} + DD^{t})^{-1}CP(t)A^{T} + BB^{T}$$
  
=  $P(t) - P(t)(P(t) + \sigma^{2}I)^{-1}P(t) + I$   
 $\hat{x}(t+1) = A\hat{x}(t) + K(t)(y(t) - C\hat{x}(t))$   
 $= \hat{x}(t) + K(t)(y(t) - \hat{x}(t))$ 

The steady-state solution of the filter is obtained by imposing that P(t+1) = P(t):

$$P(t+1) = P(t) \Rightarrow P(t)^2 = P(t) + \sigma^2 I$$

**b:** In the scalar case we can easily compute the steady state solution:

$$p^{2} - p - \sigma^{2} = 0 \Rightarrow p = \frac{1}{2} + \sqrt{\sigma^{2} + 1/4}$$

The steady-state Kalman gain becomes:

$$K = \frac{p}{p + \sigma^2}$$

In Fig. 1 the plot of a realization of the process, of the measurement, and of the estimation is shown.

c: In Fig. 2, the time varying Kalman gain is plotted for different values of  $\sigma$ .

d: In Fig. 3 the 2-dimensional case is shown.



FIGURE 1. 1-dimensional Brownian motion, measurement and estimation with  $\sigma=0.5$ 



FIGURE 2. Kalman gain K(t) for different values of  $\sigma$ 



FIGURE 3. 2-dimensional Brownian motion, measurement and estimation.