## Formula-sheet at the exam in SF2866, SF2868, 2015

If events occur according to a Poisson process with rate  $\lambda$ , where  $\tau$  denotes the time between two consecutive events, and X(T) denotes the number of events on the time interval [0,T], then  $(\lambda T)^{\ell}$ 

then  $P(\tau \le t) = 1 - e^{-\lambda t}, \quad P(X(T) = \ell) = \frac{(\lambda T)^{\ell}}{\ell!} e^{-\lambda T}, \quad \mathbf{E}[\tau] = 1/\lambda, \quad \mathbf{E}[X(T)] = \lambda T.$ <u>Markov chain in discrete time.</u>

 $\mathbf{P} = \text{the matrix with elements } p_{ij} = P(X_{n+1} = j \mid X_n = i).$ 

 $\mathbf{p}^{(n)}$  = the row vector with components  $p_j^{(n)} = P(X_n = j)$ . Then  $\mathbf{p}^{(n+1)} = \mathbf{p}^{(n)} \mathbf{P}$ . The row vector  $\pi$  defines a stationary distribution if  $\pi = \pi \mathbf{P}$ ,  $\sum_j \pi_j = 1$  and  $\pi_j \ge 0$ .

<u>Markov chain in continuous time</u> (also called Markov process with discrete state space).  $\mathbf{P}(h)$  the matrix with elements  $p_{ij}(h) = P(X(t+h) = j \mid X(t) = i)$ .  $\mathbf{p}(t) =$  the row vector with components  $p_j(t) = P(X(t) = j)$ . Then  $\mathbf{p}(t+h) = \mathbf{p}(t)\mathbf{P}(h)$ . Assumption:  $p_{ij}(h) = q_{ij}h + o(h)$  if  $j \neq i$ , while  $p_{ii}(h) = 1 + q_{ii}h + o(h) = 1 - q_ih + o(h)$ , where  $q_i = -q_{ii} = \sum_{j\neq i} q_{ij}$ . Thus,  $\mathbf{P}(h) \approx \mathbf{I} + h \mathbf{Q}$  and  $(\mathbf{p}(t+h) - \mathbf{p}(t))/h \approx \mathbf{p}(t)\mathbf{Q}$  for small h > 0. The row vector  $\pi$  defines a stationary distribution if  $\pi \mathbf{Q} = \mathbf{0}$ ,  $\sum_j \pi_j = 1$  and  $\pi_j \geq 0$ . The system  $\pi \mathbf{Q} = \mathbf{0}$  can be written  $\sum_{i\neq j} \pi_i q_{ij} + \pi_j q_{jj} = 0$ , for all j, or  $\pi_j \sum_{k\neq j} q_{jk} = \sum_{i\neq j} \pi_i q_{ij}$  ("jumps out from state j = jumps into state j").

Some quantities and relations in queueing theory (where  $P_n$  corresponds to  $\pi_n$  above):

$$L = \sum_{n=0}^{\infty} nP_n, \quad L_q = \sum_{n=s}^{\infty} (n-s)P_n, \quad \bar{\lambda} = \sum_{n=0}^{\infty} \lambda_n P_n, \quad L = \bar{\lambda}W, \quad L_q = \bar{\lambda}W_q$$
  
$$M/M/1: \quad \rho = \lambda/\mu < 1, \quad P_0 = 1-\rho, \quad P_n = \rho^n P_0, \quad L = \frac{\rho}{1-\rho}.$$
  
$$M/M/2: \quad \lambda_n = \lambda \text{ for } n \ge 0, \quad \mu_1 = \mu, \quad \mu_n = 2\mu \text{ for } n \ge 2, \quad \rho = \lambda/(2\mu) < 1,$$
  
$$P_0 = \frac{1-\rho}{1+\rho}, \quad P_n = 2\rho^n P_0 \text{ for } n \ge 1, \quad L = \frac{2\rho}{1-\rho^2}.$$

Jackson queueing networks.

Calculate  $\lambda_1, \ldots, \lambda_m$  from  $\lambda_j = a_j + \sum_i \lambda_i p_{ij}$ . Check  $\lambda_j < s_j \mu_j$ . Analyze each service facility to obtain  $P(N_j = n_j)$ . Then  $P(N_1 = n_1, \ldots, N_m = n_m) = \prod_j P(N_j = n_j)$ .  $W_1, \ldots, W_m$  can be obtained from  $W_i = V_i + \sum_j p_{ij} W_j$ , where  $V_i = L_i / \lambda_i$ .

Some deterministic inventory models.

EOQ with shortage not permitted: Minimize  $\frac{Kd}{Q} + cd + \frac{hQ}{2}$ .  $C_i = \min_j \{C_i^{(j)} \mid i \le j \le n\}$ , where  $C_i^{(j)} = C_{j+1} + K + h(r_{i+1} + 2r_{i+2} + \dots + (j-i)r_j)$ . Some stochastic inventory models.  $\overline{C(S)} = cS + p \operatorname{E}[(\xi - S)^+] + h \operatorname{E}[(S - \xi)^+]$ . If  $\xi$  is a continuous non-negative random variable then  $\operatorname{E}[(\xi - S)^+] = \int_S^{\infty} (t - S)f_{\xi}(t)dt$ ,  $\operatorname{E}[(S - \xi)^+] = \int_0^S (S - t)f_{\xi}(t)dt$ , and  $C'(S) = c + p \left(F_{\xi}(S) - 1\right) + hF_{\xi}(S)$ . If  $\xi$  is a non-negative integer-valued random variable then S is integer and  $\operatorname{E}[(\xi - S)^+] = \sum_{j=S}^{\infty} (j - S)p_{\xi}(j)$ ,  $\operatorname{E}[(S - \xi)^+] = \sum_{j=0}^S (S - j)p_{\xi}(j)$ , and  $C(S + 1) - C(S) = c + p \left(F_{\xi}(S) - 1\right) + hF_{\xi}(S)$ . Marginal allocation for generating efficient solutions to the pair (f, g), where f and g are integer-convex separable functions, f decreasing and g increasing in the non-negative integer variables  $x_1, \ldots, x_n$ . Generate a table in which the j:th column contains the quotients  $-\Delta f_j(0)/\Delta g_j(0), -\Delta f_j(1)/\Delta g_j(1), -\Delta f_j(2)/\Delta g_j(2), \ldots$ Let all the quotients in the table be uncanceled.

Initiate the variables to their smallest feasible values and repeat the following:

Let  $\ell$  be the number of the column with the largest uncanceled quotient.

Cancel this quotient, and increase the  $\ell$ :th variable  $x_{\ell}$  by one.

<u>Finite horizon MDP recursion</u> (discounting if  $0 < \alpha < 1$ , no discounting if  $\alpha = 1$ ):

$$V_i^{(n)} = \min_k \{ C_{ik} + \alpha \sum_j p_{ij}(k) V_j^{(n-1)} \}$$
 (backward time).

Policy improvement algorithm for MDP without discounting:

1. For a given policy, calculate  $v_0, \ldots, v_M$  and g from  $v_M = 0$  and  $g + v_i = C_{i,d_i} + \sum_j p_{ij}(d_i)v_j$ .

2. The current policy is optimal if  $g + v_i = \min_k \{ C_{ik} + \sum_j p_{ij}(k)v_j \}$ . Otherwise, define a new policy by letting  $d_i =$  a minimizing k above. Then go to 1.

Policy improvement algorithm for MDP with discounting:

1. For a given policy, calculate  $V_0, \ldots, V_M$  from  $V_i = C_{i,d_i} + \alpha \sum_j p_{ij}(d_i)V_j$ .

2. The current policy is optimal if  $V_i = \min_k \{ C_{ik} + \alpha \sum_j p_{ij}(k)V_j \}$ . Otherwise, define a new policy by letting  $d_i =$  a minimizing k above. Then go to 1.

Note: No calculator at the exam!