



KTH Teknikvetenskap

Harald Lang

Tentamensskrivning, 12/03-2009, kl 14.00–19.00.  
SF2951 Ekonometri  
Short answers

1. The “normal equations” says that  $\sum_1^n x_i \hat{e}_i = 0$ . Note that  $\hat{y}_i = x_i' \hat{\beta} = \hat{\beta}' x_i$ . We get

$$\begin{aligned} \sum_1^n y_i^2 &= \sum_1^n (\hat{\beta}' x_i + \hat{e}_i)^2 = \sum_1^n (\hat{\beta}' x_i)^2 + 2 \sum_1^n \hat{\beta}' x_i \hat{e}_i + \sum_1^n \hat{e}_i^2 \\ &= \sum_1^n \hat{y}_i^2 + 0 + \sum_1^n \hat{e}_i^2 \end{aligned}$$

*Q.E.D.*

2. Let  $y_i =$  dummy for “not repaid”,  $x_{1,i} =$  dummy for “record of non-payment during the last two years”,  $x_{2,i} =$  dummy for “record of non-payment during the period two to five years back in time” and  $x_{3,i} =$  “interest rate of the loan”. Here  $i$  refers to borrower number  $i$ . We model the probability  $p(x_1, x_2, x_3)$  as a “Logit”:

$$p(x_1, x_2, x_3) = \frac{\exp(\beta_0 + x_1 \beta_1 + x_2 \beta_2 + x_3 \beta_3)}{1 + \exp(\beta_0 + x_1 \beta_1 + x_2 \beta_2 + x_3 \beta_3)}$$

and either estimate this equation by ML, or the equation

$$y = \frac{\exp(\beta_0 + x_1 \beta_1 + x_2 \beta_2 + x_3 \beta_3)}{1 + \exp(\beta_0 + x_1 \beta_1 + x_2 \beta_2 + x_3 \beta_3)} + error$$

by GNLS (known covariance matrix).

- 3a. There are more than one possible test, but a Wald test seems most natural.  
b. If the null is true, then the Wald test says that

$$\frac{1.48^2}{0.45} + \frac{1.36^2}{0.38} \approx 9.73$$

is an observation of a  $\chi^2(2)$ -variable. Since this exceeds 9.21, we reject the null at the 1% level.

- 4a. The interpretation of  $x' \beta$  is  $x' \beta = \text{median}(y | x)$ .  
b. Possible reasons include: we *want* to estimate the median rather than the expected value of  $y$ , the estimation is more robust to outliers, invariance to monotonic transformations.
- 5a. Solve for the reduced form. We get

$$y_2 = \frac{1}{1 - \alpha_2 \beta_2} ([\beta_0 + \beta_2 \alpha_0] + \beta_2 \alpha_1 x_1 + \beta_1 x_2 + [e_2 + \beta_2 e_1])$$

so we see that  $y_2$  is correlated with  $e_1$  (unless  $\beta_2 = 0$ .)

- b. Both equations can be estimated with IV (2SLS) with the exogenous variables  $x_1$  and  $x_2$  as instruments.

6. We re-estimate  $\beta$   $N$  times ( $N = 2'000$  or some such) using resampling of with replacement of  $n$  observations. Denote the new estimates  $\hat{\beta}_i^*$ ,  $i = 1 \dots, n$ . Now compute the values  $g(x_0; \hat{\beta}_i^*)$ ,  $i = 1 \dots, n$ , and choose  $a$  and  $b$  such that 0.5% of the values  $g(x_0; \hat{\beta}_i^*)$  falls above  $b$ , and 0.5% of the values  $g(x_0; \hat{\beta}_i^*)$  falls below  $a$ . Now an approximate 1% confidence interval for  $g(x_0; \beta)$  is

$$2g(x_0; \hat{\beta}) - b \leq g(x_0; \beta) \leq 2g(x_0; \hat{\beta}) - a.$$

Here  $\hat{\beta}$  is the original estimate. Similarly for any other level of confidence.