Estimation of Loss Given Default for Low Default Portfolios

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ABSTRACT

The Basel framework allows banks to assess their credit risk by using their own estimates of Loss Given Default (LGD). However, for a Low Default Portfolio (LDP), estimating LGD is difficult due to shortage of default data. This study evaluates different LGD estimation approaches in an LDP setting by using pooled industry data obtained from a subset of the PECDC LGD database. Based on the characteristics of a LDP a Workout LGD approach is suggested. Six estimation techniques, including OLS regression, Ridge regression, two techniques combining logistic regressions with OLS regressions and two tree models, are tested. All tested models give similar error levels when tested against the data but the tree models might produce rather different estimates for specific exposures compared to the other models. Using historical averages yield worse results than the tested models within and out of sample but are not considerably worse out of time.
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LIST OF ABBREVIATIONS

A-IRB       Advanced Internal Rating Based
F-IRB       Foundation Internal Rating Based
EAD         Exposure At Default
EL          Expected Loss
IRB         Internal Rating Based
LDP         Low Default Portfolio
LGD         Loss Given Default
PD          Probability of Default
PIT         Point In Time
SDR         Standard Deviation Reduction
SME         Small and Medium Enterprises
TTC         Through The Cycle
UL          Unexpected Loss
VaR         Value at Risk

Models
FT          F-Tree \textit{(introduced in this paper)}
LR-OLS      Logistic-OLS Regressions \textit{(introduced in this paper)}
OLS         Ordinary Least Squares regression
RiR         Ridge Regression
RT          Regression Tree
TLR-OLS     Trimmed Logistic-OLS Regression \textit{(introduced in this paper)}
Hist        Historical average

Model evaluation methods
MAE         Mean Absolute Error
ME          Mean Error
R2          R-squared value
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1 INTRODUCTION

A sound and stable financial system is an essential part of a growing and prosperous society. In contrast, a financial crisis can severely damage the economic output of a society for an extended period of time, as seen in recent years. In order to reduce the risk of future crises, banks and other financial institutions are regulated. The outlines of banking regulations are set in the Basel accords, where banks are deemed to have a certain amount of capital in order to cover potential future losses.

The risks facing banks are multifaceted and therefore divided into several parts, the main ones being credit risk, market risk and operational risk. In order to assess the credit risk, certain risk parameters must be estimated. These include Probability of Default (PD), Loss Given Default (LGD) and Exposure At Default (EAD). The Basel framework defines three possible approaches for assessing the credit risk exposure. The Standardized approach, the Foundation Internal Rating Based approach (F-IRB) and the Advanced Internal Rating Based approach (A-IRB). Differing to the standardized and F-IRB approach, the A-IRB approach requires the bank to use its own estimates of LGD and EAD in addition to PD. In the standardized and F-IRB approach the LGD parameter is given by the regulators. This thesis focuses on the estimation of LGD. Historically, a lot of focus has been devoted to the estimation of PD while LGD has received less attention and has sometimes been treated as constant. Das and Hanouna (2008) note that using constant loss estimates might be misleading since losses experience large variation in reality. According to Moody’s (2005) average recovery rates, defined as 1- LGD, can vary between 8% and 74% depending on the year and the bond type. For a sophisticated risk management, LGD clearly needs to be assessed in more detail.

The estimation of LGD is preferably conducted by using historical loss data, but for certain portfolios and counterparties, there is a shortage of such data due to the high quality of the assets and the low number of historical defaults. Portfolios of this kind are often referred to as Low Default Portfolios (LDPs). LDPs include portfolios with exposures to e.g. banks, sovereigns and highly rated corporations. The most extreme examples include covered bonds with very few cases where a loss has occurred. For LDPs the estimation of credit risk parameters like LGD is problematic and this has led to questions from the industry concerning how to handle these portfolios.

The purpose of this thesis is to study quantitative models for estimation of LGD and empirically evaluate how these models work on LDPs in order to find a model that can be used in practice. For the model to be useful in practice it must produce reasonable and justifiable values despite little default data. While the models are based solely on quantitative factors also qualitative considerations are taken into account when the models are constructed.

The outcome of the study will constitute of two parts, the first being an overview of the academic progress in this area and the second an evaluation of models from a practical perspective. The direct beneficiaries of the thesis are banks and financial institutions with the need to assess their credit risk exposure.

The benefit of a better credit risk assessment is twofold. First it gives banks a better control over the risks they are facing and can be a support for business decisions. Secondly, internal models typically results in lower risk measures and thereby lower capital requirements. Since capital is costly, this is a direct benefit for a bank. On a more general level, society as a whole benefits from sound financial institutions with good credit risk assessments.

The remainder of this report is structured as follows. Chapter 2 presents a theoretical background to LGD and LDPs. In Chapter 3 the models used in this study are presented. Chapter 4 gives an overview of the data used for this study while
Chapter 5 presents the evaluation methods used. In Chapter 6 the results are presented and Chapter 7 contains concluding remarks.
2 THEORETICAL BACKGROUND

2.1 DEFINITION OF DEFAULT

The event of default can be defined in many ways and it differs for different purposes and disciplines. A default from a legal point of view is not necessarily the same as a default from an economic point of view. For the purpose of this paper it is important that the definition

1) Complies with relevant regulations
2) Is consistent with the definition used to estimate the probability of default parameter (PD)

For these reasons we use the definition communicated in the Basel framework (BIS, 2006, § 452)

A default is considered to have occurred with regard to a particular obligor when either or both of the two following events have taken place.

- The bank considers that the obligor is unlikely to pay its credit obligations to the banking group in full, without recourse by the bank to actions such as realising security (if held).
- The obligor is past due more than 90 days on any material credit obligation to the banking group. Overdrafts will be considered as being past due once the customer has breached an advised limit or been advised of a limit smaller than current outstandings.

When this definition is used a default occurs when the obligor is unlikely to pay or is late with a payment. As the obligor does not necessarily have to miss a payment for a default to occur it is not impossible that the obligor will actually pay the full amount on time and the loss be equal to zero.

2.2 DEFINITION OF LOSS GIVEN DEFAULT (LGD)

LGD is the economic loss occurring when an obligor defaults. It is expressed in percentage of the defaulted amount. The defaulted amount is equal to the principal amount and overdue interest payments. This is consistent with the finding of Andritzky (2005) that the core legal claim consists of the nominal value and is often referred to as the recovery of face value assumption. In accordance with the Basel framework the LGD estimates are based on the economic loss and include workout costs arising from collecting the exposure (BIS, 2006, § 460). These costs are included in the modelled LGD since it is believed that they are affected by the exposure type. Complicated bankruptcy processes might imply a larger use of external lawyers for instance. Since workout costs are included, the LGD value can be larger than one. However, since this is not believed to occur frequently, LGD is assumed to be equal to or lower than one and LGD observations and estimates are therefore capped at one.

The cash flows received during the workout process are adjusted with a discount rate equal to the risk free interest rate at the time of default. The use of the risk free interest rate for calculating the present value of the cash flows is a simplification. In practice regulators require the discount rate to include an appropriate risk premium (BIS 2005a). However, there is no consensus among practitioners and banking supervisors regarding what discount rate to apply for different cases (Brady et al 2006). Since the addition of a risk premium is not believed to affect the choice of a model it is not considered in this study.
2.3 Definition of Low Default Portfolio (LDP)

There is no exact definition of a LDP accepted by the industry (BIS, 2005b). Instead, the Basel Committee Accord Implementation Group’s Validation Subgroup points out that a bank’s portfolio is not either low default or non-low default but that there is “a continuum between these extremes” and notes that “a portfolio is closer to the LDP end of this continuum when a bank’s internal data systems include fewer loss events, which presents challenges for risk quantification and validation” (BIS, 2005b § 1).

The International Swaps and Derivatives Association (ISDA) notes in a joint working paper (2005) that examples of LDPs include e.g. exposures to sovereigns, banks, large corporates and repurchase agreements (ISDA, 2005).

2.4 The Mathematical Nature of Loss Given Default (LGD)

LGD is the loss conditional upon the event of default. The parameter depends therefore crucially on the default definition. While the definition varies for different purposes and disciplines this paper use the definition from the Basel accords, see Section 2.1.

The definitions of LGD and the default event have two important implications.

1) Only losses of cash flows (principal and overdue interest payments and workout costs) are considered. Losses purely due to market movements or a changed market price of the underlying are not considered.
2) A loss cannot occur without the event of default since a loss of a principal or interest payment necessarily leads to the state of default.

Considering a probability space \( \Omega, \mathcal{F}, P \) two random variables can be defined:

\[
D = \begin{cases} 
1 & \text{if default occurs} \\
0 & \text{otherwise} 
\end{cases} \\
L \in [0,1] \text{ representing the loss of an exposure.}
\]

Loss given default can then be defined as the random variable \( \text{LGD} \equiv L|D = 1 \).

Furthermore three important outcomes can be identified:

\[
A: D = 0; L = 0 \\
B: D = 1; L = 0 \\
C: D = 1; 0 < L \leq 1.
\]

It should be noted that due to the definitions of loss and default the following holds for the loss variable \( L \): \( P(L = 0|D = 0) = 1 \).

This leaves the case when \( D = 1 \).

Equation (1) highlights the importance of a consistent default definition for the PD and LGD parameters since the PD parameter is in fact part of the LGD random variable. PD is defined as \( P(D=1) \) in
\[ P(LGD = a) \equiv P(L = a | D = 1) = \frac{P(L = a \cap D = 1)}{P(D = 1)} = \frac{P(L = a \cap D = 1)}{PD} \quad (1) \]

The input to the credit risk reporting will be the expected value of the random variable LGD, \( \bar{LGD} = E[LGD] \).

Contrary to the mathematics derived in this paper regulators require the LGD parameter to be a “downturn LGD” reflecting the LGD in a downturn environment and not the expected value of the random variable. The Basel committee suggests two ways of deriving a downturn LGD. One could either use a mapping function to extrapolate the normal LGD or one could provide a separate downturn LGD estimate.

### 2.5 The Low Default Portfolio (LDP) Problem

LDPs are not mentioned explicitly in the Basel II framework (Kofman, 2011) but portfolios of this kind have raised concerns from the industry, fearing that these portfolios might end up excluded from IRB-treatment. The Basel committee has acknowledged these concerns and stated in a newsletter in 2005 that LDPs should not automatically be excluded from IRB-treatment. The committee suggested some remedies for the shortage of data including pooling of loss data and combination of portfolios with similar characteristics (BIS, 2005b). However, no models or solutions are proposed if the problem persists.

Most studies concerning the estimation problems for LDPs are focusing on the problem of estimating the probability of default parameter even though the LDP problem may be even more severe for LGD estimation. ISDA (2005) notes that there sometimes might exist a sufficient amount of defaults to estimate PD but too few observations to estimate LGD. Kofman (2011) suggests that LDPs may be extended by “near defaults” (called “quasi-defaults”) with a similar financial profile in order to overcome the issue of too little data. A near default might be identified by a high risk profile or a low/downgraded credit rating.

While this might be helpful for the estimation of PD it seems unlikely that it would improve the LGD estimation since these observations per definition has a loss of zero. Other proposed ways of estimating PD includes for instance the use of migration matrices and Markov Chains. In these models migration rates between rating grades are used in order to assess the likelihood of a highly rated asset to be downgraded and eventually defaulted (Schuermann and Hanson, 2004). This approach requires default data for lower rating grades and might therefore not be applicable when there are a low number of defaults in the whole portfolio. Furthermore, even if there is some evidence that rating grades do affect the recovery rates (Altman et al., 2004) it is hard to see how to extend the data for LGD in the same way as for PD estimations.

Pluto and Tachse (2005) present a model using the “most prudent estimation” principle that does not share the requirement of defaults for lower rating grades. Their model assumes that the initial ranking is correct and an asset with a lower rating cannot have a lower probability of default than an asset with a higher rating. They derive confidence intervals for the probability of default for every rating grade based on the number of defaults and PD parameters are estimated as the upper confidence bounds. This guarantees that the differences between the credit ratings are preserved. A disadvantage of the model is the prerequisite of a ranking system. Furthermore, while the use of ratings might be sensible for estimating PD this might not be appropriate for the estimation of LGD.

While the LDP problem affects all components of the expected loss (PD, LGD and EAD) most focus has so far been on the estimation of PD. Apart from pooling of data, suggested remedies are unfortunately not applicable on LGD estimation.
2.6 POINT IN TIME, THROUGH THE CYCLE AND DOWNTURN ESTIMATES

There are two main approaches used when defining parameters for risk estimation, Point in Time (PIT) and Through the Cycle (TTC). A PIT estimate is constructed to capture the risk at every point in time while a TTC estimate is constructed to capture the average risk through a business cycle. This study aims to provide a TTC estimate and discuss the inclusion of macroeconomic variables only as a way of generating the downturn estimate required by regulators.

PECDC (2013a) notes that observations of defaulted banks and financial institutions are typically connected to crises and LGD estimates for banks and financial institutions are therefore already associated with downturns in the financial markets. It can therefore be questioned to what degree separate downturn LGD estimates have to be produced for this particular type of exposures. Looking at the sample in this study, more than half of the data is associated with downturn periods.

2.7 LGD ESTIMATION APPROACHES

As CEBS (2006) does not suggest a specific model or approach for LGD estimation but merely concludes that "supervisors do not require any specific technique for LGD estimation" (p. 72) the choice of method is an important part of LGD estimation. CEBS (2006) lists four main techniques for quantitative estimation of LGD. These are Workout LGD, Market LGD, Implied Market LGD and Implied Historical LGD. Workout LGD is estimated from cash flows from the workout process and Market LGD estimates are derived from the prices of defaulted bonds. Implied Market LGD is instead derived from the prices of non-defaulted bonds or derivatives on said bonds and could thus be used to estimate the LGD without the security actually defaulting. However, CEBS (2006) points out that Market LGD and Implied Market LGD only can be used in limited circumstances and probably not for the main part of a loan portfolio. Implied Historical LGD is only allowed for retail exposures and hence cannot be applied for the purpose in this study.

WORKOUT AND MARKET LGD

The choice between Workout and Market LGD is essentially the choice of when to observe the recovery rate. Workout LGD estimations use ultimate recoveries from the workout process while Market LGD estimations use the trading prices at some time after default. The advantage of the Workout LGD procedure is that it uses the true values of recoveries while Market LGD, affected by the supply and demand, risk aversion and liquidity of the post-default market, has been found to systematically underestimate the true recovery rates (Renault and Scaillet, 2004). A Market LGD estimation also requires that the defaulted security is actually traded on a public market. However, using Market LGD makes it possible to include defaults that occurred recently in the model. As a workout processes can last several years (Araten, 2004) this is a considerable advantage when data is scarce. Table 1 lists possible approaches for the estimation of LGD using observed Workout or Market LGD.
**Estimation techniques**

**Parametric:**
- Ordinary least squares regression (OLS) (e.g. Qi and Zhao, 2012)
- Ridge regression (RiR) (Loterman, 2012)
- Fractional response regression (Bastos, 2010)
- Tobit model (Calabrese, 2012)
- Decision tree model (Logistic-OLS Regressions model (LR-OLS)) (Calabrese, 2012)

**Transformation regressions:**
- Inverse Gaussian regression (Qi and Zhao, 2012)
- Inverse Gaussian regression with beta transformation (Qi and Zhao, 2012)
- Box-Cox transformation/OLS (Loterman, 2012)
- Beta transformation/OLS (Loterman, 2012)
- Fractional logit transformation & Log transform (Bellotti and Crook, 2008)

**Non-parametric:**
- Regression tree (RT) (Bastos, 2010)
- Neural networks (Qi and Zhao, 2012)
- Multivariate adaptive regression spline (Loterman, 2012)
- Least squares support vector machine (Loterman, 2012)

**Semi-parametric:**
- Joint Beta Additive Model (Calabrese, 2012)

**Table 1. Workout and Market LGD estimation techniques**

**IMPLIED MARKET LGD**

In order to estimate LGD with market data of non-defaulted securities the theory of pricing debt is used. In theory the price of defaultable debt is governed by the perceived PD and LGD. Assuming a constant default rate, PD, the value at time \( t = 0 \) of a defaultable coupon paying bond could, in theory, be expressed as

\[
V_0 = \sum_{t=1}^{N} [(1 - PD)^t \cdot C_t] + (1 - PD)^N \cdot d_N \cdot FV + \sum_{t=1}^{N} [PD^t \cdot d_t \cdot C_t \cdot (1 - LGD)] + PD^N \cdot d_N \cdot FV \cdot (1 - LGD),
\]

where \( FV \) is the face value paid at time \( N \) in case of no default, \( C_t \) is the coupon paid at time \( t \) in case of no default, PD is the default rate per time period, \( d_t \) is the discount factor from time \( t \) to 0 and LGD is the loss encountered in case of default. However the price is also, just as in the Market LGD case, affected by liquidity, risk aversion and supply and demand aspects which are not accommodated in the model.

To be able to extract LGD from this equation the problem of separating the effect of PD and LGD has to be solved, the so called identification problem. Studies propose different solutions to this. Schläfer and Uhrig-Homburg (2014) and Unal et al. (2003) suggest the use of debt with different seniority in the same firm. The only difference in prices between these debts should be the effect of implied LGD as the probability of default is the same for all securities. Other similar approaches using
credit default swaps and equity (Das and Hanuona, 2008) as well as credit default swaps and bond spreads (Andritzky, 2006) and digital default swaps and bonds (Song, 2008) are also proposed in the literature.

<table>
<thead>
<tr>
<th>Identification problem remedies:</th>
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<tr>
<td>Credit default swaps &amp; Digital default swaps (Berd, 2004)</td>
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<tr>
<td>Credit default swaps &amp; Equity (Das and Hanuona, 2008)</td>
</tr>
<tr>
<td>Junior vs Senior debt (Schläfer and Uhrig-Homburg, 2014) (Unal et al., 2003)</td>
</tr>
<tr>
<td>Credit default swaps vs Bond spreads (Andritzky, 2006)</td>
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**2.8 LGD ESTIMATION APPROACHES FOR LDPS**

As previously noted, the LDP remedies suggested for the estimation of PD, e.g. inclusion of “near-defaults” or using migration matrices (see Section 2.5) are in general not possible to use for the estimation of LGD. The one remedy which has a substantial effect is to pool data and base the estimation not only on banks’ own LGD observations.

Since the number of historical LGD observations is low, also after pooling, it could be tempting to use an implied market estimation approach. However, this approach experiences several difficulties, the first one being the identification problem mentioned earlier. However, also with the identification problem solved some problems persist. In practice the price of defaultable debt is also affected by the debt’s liquidity and the risk aversion of the market. This problem is even more severe for a low default portfolio where the effects of PD and LGD on the prices are small. Christensen (2005) concludes that “for firms of very high credit quality (A-rated companies and above) the default intensity is so low that it is close to impossible to measure the risk contribution from the stochastic recovery”. Andritzky (2005) states that in order to be able to determine the recovery the bonds should contain a “considerable default risk”. Otherwise the effect of the recovery rate is too small to be measured accurately. Since an LDP typically consists of exposures with a very small default probability, implied market LGD estimation approaches are inappropriate for this type of portfolio.

Workout LGD based on pooled data would be possible to apply on a LDP if enough data is available. Also a Market LGD estimation approach could be justified to use on a LDP from a theoretical viewpoint since it increases the possible sample with defaults that occurred recently. However not all securities that might be in a LDP are traded on a market. Furthermore, the market prices are based on the market’s estimation of the future recovery and the prices are affected by liquidity aspects and risk aversion.

**2.9 WHAT AFFECTS LGD? (RISK DRIVERS)**

If a Workout or Market estimation approach is used, the set of explanatory variables to include in the models is an important aspect of the estimation problem. The theoretical progress regarding risk drivers for LGD levels are summarized below. None of the studies referred to have however used a sample with financial institutions.

**MACROECONOMIC ENVIRONMENT**

Several studies (e.g. Schuemann, 2004 and Khieu et al., 2012) note that recoveries are lower in recessions. The exact magnitude is uncertain but Frye (2000) indicates that bond and loan recoveries could decrease by 20% during recessions.
Since bankruptcy processes can last several years (Araten, 2004) an important aspect when looking at the macroeconomic environment is the time lag between the event of default and the bankruptcy process where the firm's assets are sold. While the macroeconomic conditions at the event of default might influence the probability of default what probably influences the LGD is the macroeconomic environment during the bankruptcy process. This could potentially explain why some studies (e.g. Schuermann, 2004) do not find a clear relation between the macroeconomic environment and the LGD levels. Another reason could be the proxy for the macroeconomic conditions used. Most studies (e.g. Khieu et al., 2012) use GDP growth as such a proxy but others (e.g. Unal et al., 2001) have proposed to use the interest rate. An interest rate as a proxy has the benefit of being publically available at every point in time and not reported subsequently as GDP. It could also be argued that interest rates are to some extent forward looking in a way that GDP growth figures are not and hence would capture the conditions of the economy during the bankruptcy process in a better way. In addition to determining what measures to use as a proxy for the state of the economy it needs to be decided what geographic division to use. While a small enterprise might be mostly affected by the state of the domestic economy a larger multinational enterprise might be more affected by the state of the world economy than by the state of the economy where its headquarter is incorporated.

Some studies report a positive correlation between the probability of default and the realized levels of LGD (Altman et al., 2004). However, this is probably linked to the same causality, the macroeconomic environment. During a recession, many companies default and the realized LGD levels are higher. It is important to note that this is different from saying that a company with a high probability of default is likely to have a high LGD.

**SENIORITY & FACILITY TYPE**

In theory the Absolute Priority Rule implies that higher seniorities of debt should be repaid in full before lower ranked debt receives anything. While this could be violated in practice (see e.g. Weiss, 1990), securities with a higher seniority should experience lower LGD levels in general. This is confirmed by Schuermann (2004) who suggests that seniority is the most important factor in determining LGD levels. The seniority of the debt is closely linked to the facility type. Loans for instance, typically experience lower LGD levels than bonds since they typically have a higher seniority (Schuermann, 2004). Few academic studies use other facility types than bonds and loans but Khieu et al. (2012) find a significant difference in LGD levels between term loans and revolving credits.

**COLLATERAL**

Many studies conclude that the degree of securitization is one of the most important factors for determining LGD. According to Dermine and Carvalho (2006) the type of the collateral is also an important aspect. Dermine and Carvalho (2006) distinguish between real estate, financial and physical collateral but all types of collateral show a positive correlation with recovery rates.

**GEOGRAPHIC REGION**

The country of the borrower is widely used as a risk driver in credit risk modelling. Since legal differences in the bankruptcy process may affect the LGD, the geographical region has been used by e.g. Gupton (2005) and Araten (2004) in LGD modelling.
INDUSTRY
According to Schuermann (2004) the industry of the obligor affects the LGD levels. This is especially important in the case of industries with a lot of tangible assets, like utilities, which experience lower LGD levels than industries with low levels of tangible assets like for instance the service industry. This is due to the fact that tangible assets in contrast to intellectual ones tend to maintain their value in a bankruptcy process.

SIZE OF FIRM & EXPOSURE
Some studies propose that the size of the obligor would affect LGD levels. Dermine and Carvalho (2006) argue that banks would be more reluctant to default a large loan because of “spill-over effects” on other loans to the bank, and that large loans actually defaulting will be in worse shape. They also empirically find a positive effect of the loan size on bank loans’ realized LGD. In contrast, Schuermann (2004) reports no effect of loan size on the LGD of bank loans. However, the relationships may also be rather different between banks and SMEs compared to between financial institutions.

INDUSTRY DISTRESS
When assets are industry specific a potential buyer is likely to be a competitor and the state of the industry is then an important factor for determining LGD levels in addition to the general macroeconomic environment (Acharya et al., 2007).

LEVERAGE
The firm’s leverage has often been considered to influence the LGD levels. A high level of leverage means that the assets of the firm needs to be shared among more debt holders which should influence the LGD levels positively (Schläfer and Uhrig-Homburg, 2014). Furthermore, it has been suggested that a high leverage ratio may be associated with a more scattered debt ownership implying longer and more complicated bankruptcy processes, also increasing the LGD levels (Acharya et al 2007). However, it has also been proposed that a high leverage may influence the LGD levels negatively since a high leverage may be followed by an increased monitoring activity (Khieu et al 2012). However, this argument is probably easier to justify for e.g. SMEs than for financial institutions.

GUARANTEE
Several studies report the effect of guarantees on LGD levels, see e.g. Qi (2011) and Dermine (2005). A guarantee should in theory result in lower realized LGD levels. However, as pointed out in the study by Dermine (2005), guarantees (and collaterals) may also be an indication of a greater risk since it is usually not requested from “good clients”.

UTILIZATION RATE
Since firms sometimes maximize their credit lines in order to avoid default the utilization rate could be a predictor for the LGD level. It is however doubtful how functional such a variable would be in practice since the utilization rate probably soars just before the event of default while it is moderate at some earlier point in time. This would be problematic when trying to estimate LGD levels in practice.
DEFAULT YEAR

Before the release of the Basel II default definitions (see Section 2.1), most credit risk models used the event of bankruptcy as the default definition (Hayden, 2003). Since the Basel II definition is much stricter, more situations classify as defaults if these rules are applied. The fact that there is not one single definition of default and the lack of consistency of this definition through time can be problematic when modelling credit risk since the definition might not be consistent throughout the sample. Because of this, estimated LGD levels might be affected by when the majority of the observations were reported.
3 MODELS

Based on the characteristics of the LGD estimation approaches, a Workout LGD estimation approach has been adopted in this study and six different estimation techniques are tested. The reasons for choosing the following estimation techniques are simplicity, degree of computational intensiveness and the possibility to apply on a relatively small sample of LGD observations. All models are purely quantitative and based on a number of quantitative risk drivers. The models are however constructed also based on qualitative considerations when it comes to parameter selection, see Section 6.1. In practice it is however not uncommon to include qualitative risk drivers of an exposure, such as management, ownership and the risk culture of the business, in credit risk modelling (ISDA, 2005).

Out of the many possible estimation techniques based on the Workout LGD approach the ones outlined below have been chosen for empirical testing. Two new estimation techniques, the Trimmed Logistic-OLS Regressions model and the F-Tree model are also included. The first four models are all based to some extent on a linear regression while the two final models are based on tree structures.

Tree models are non-linear and non-parametric models and are therefore more flexible in capturing data characteristics. Unfortunately, these kinds of models are also much more prone to overfitting. The reason for this is that a larger tree results in better fit but possibly worse predictions out of sample. It could be noted that the extreme case of just a single observation in every leaf will result in a perfect fit but probably rather poor out of sample predictions.

When constructing a tree model qualitative considerations regarding which splits to perform could also be included. Qualitative considerations could include e.g. which splits make sense from a theoretical viewpoint as well as the number of observations necessary in each node or in the Regression tree case how large the reduction must be in order to be included in the model.

ORDINARY LEAST SQUARES REGRESSION (OLS)

Ordinary least squares regression is the most commonly used regression technique. It is proposed and tested in many academic studies, although not in a LDP setting, see e.g. Qi and Zhao (2012) and Loterman (2012). OLS minimizes the sum of squared errors and is the best linear unbiased estimator (Lang, 2012). In order to capture non-linear effects the parameters can be squared and logged etc. but this is not considered in this study. Several studies (e.g. Khieu et al., 2012; Qi and Zhao, 2012) suggest however that the OLS method is ill-suited for LGD data due to the bimodal distribution and the bounds at 0 and 1.

Because of the boundary conditions the LGD estimates are truncated to the [0,1] interval afterwards. Since an OLS regression minimizes squared errors the resulting estimates will be more conservative than if the absolute errors were minimized (Bellotti and Crook, 2008). This is reasonable for a LDP since the Basel accords encourage conservative estimates when less data is available (BIS, 2006, § 411).

RIDGE REGRESSION (RI-R)

Loterman (2012) proposes the use of a Ridge Regression (or Tikhonov–Miller method, Phillips–Twomey method, constrained linear inversion, linear regularization) for modelling LGD. It is similar to an OLS regression but tries to regularize ill-posed problems. A Ridge regression is therefore less sensitive to correlated independent variables. Since the sample in this study is small and the same default can be reported more than once by different creditors with similar exposures it is not unlikely that some of the variables in the sample are correlated. In the same way as for the OLS model, the LGD estimates are truncated to the [0,1] interval because of the boundary conditions.
RiR seeks to minimize the expression \( \|Ax - b\|^2 + \|Tx\|^2 \), including a chosen Tikhonov matrix \( T \) in addition to the Euclidean norm. \( T \) is often set to the identity matrix. An explicit solution to this optimization problem is

\[
\hat{x} = (A^*A + T^*T)^{-1}A^*b,
\]

where \( A^* \) and \( T^* \) represent the transpose of matrix \( A \) respectively matrix \( T \). The effect of the regularization may be varied with the scale of \( T \), where \( T = 0 \) gives the unregularized least squares solution.

**LOGISTIC-OLS REGRESSIONS (LR-OLS)**

A model including logistic regressions has been proposed by e.g. Bellotti and Crook (2012). It is based on the idea that special circumstances could lead to full or none recovery of the exposure. In order to capture this two separate logistic regressions for the special cases \( \text{LGD}=0 \) and \( \text{LGD}=1 \) are performed in addition to an OLS regression for the case \( 0 < \text{LGD} < 1 \). Logistic regressions are appropriate to use when the dependent variable is binary and are in this case used to estimate the probability of \( \text{LGD}=0 \) and \( \text{LGD}=1 \). To estimate the parameters a maximum likelihood estimation is performed for each logistic regression. Bellotti and Crook (2012) and Calabrese (2012) call this model the “Decision Tree model”, however, in order to avoid confusion the model is here called Logistic-OLS Regressions model since the model differs in nature from the models called tree models in this study.

Following the approach outlined by Bellotti and Crook (2012) the estimated LGD is calculated as

\[
E[\text{LGD}] = P(\text{LGD} = 0) \cdot 0 + (1 - P(\text{LGD} = 0)) \cdot \left( 1 - \frac{1}{(1 - P(\text{LGD} = 0)) \cdot \text{LGD}_{\text{OLS}}} \right)
\]

\[
= (1 - P(\text{LGD} = 0)) \cdot \left( 1 - \frac{1}{(1 - P(\text{LGD} = 0)) \cdot \text{LGD}_{\text{OLS}}} \right) \cdot \text{LGD}_{\text{OLS}}.
\]

where \( \text{LGD}_{\text{OLS}} \) is estimated from the OLS regression and \( P(\text{LGD} = 0) \) and \( P(\text{LGD} = 1|\text{LGD} \neq 0) \) are estimated from logistic regressions.

**TRIMMED LOGISTIC-OLS REGRESSIONS (TLR-OLS)**

As an alternative to the Logistic-OLS Regressions model above we suggest a model based on the idea that while the case of no recovery might not be fundamentally different from other recovery levels the case of full recovery might bear special characteristics. This method could also potentially be better suited for small datasets than the Logistic-OLS Regressions model since the sample is divided into two instead of three parts. This model has been given the name Trimmed Logistic-OLS Regression model since two of the cases in the Logistic-OLS Regressions model have been merged. To calculate a LGD estimate the expected value of LGD is calculated as

\[
E[\text{LGD}] = P(\text{LGD} = 0) \cdot 0 + P(0 < \text{LGD} \leq 1) \cdot \text{LGD}_{\text{OLS}} = (1 - P(\text{LGD} = 0)) \cdot \text{LGD}_{\text{OLS}}.
\]

where \( \text{LGD}_{\text{OLS}} \) is estimated from the OLS regression and \( P(\text{LGD} = 0) \) from a logistic regression.

**REGRESSION TREE (RT)**

Many academic studies have proposed the use of regression tree models, e.g. Qi and Zhao (2012), Bastos (2010) and Loterman (2012), for the modelling of LGD. Regression trees are non-parametric and non-linear models which are based on a greedy search algorithm splitting the dataset into smaller and smaller subsets. A greedy search algorithm is an algorithm which “always takes the best immediate, or local, solution while finding an answer” (Black, 2005). This type of algorithm will of course
not always find the optimal solution but it is much less computationally intensive than to find the globally optimal solution. In this case the algorithm searches over all possible splits in order to find the split minimizing the intra-variance of the subsets. This is repeated until a certain stopping criterion is reached. The final subsets are called leaves. Bastos (2010) proposed to measure the decrease in variance by the "standard deviation reduction" (SDR) defined as

$$SDR = \sigma(T) - \frac{\bar{T}_1}{T} \cdot \sigma(T_1) - \frac{\bar{T}_2}{T} \cdot \sigma(T_2)$$

where $T$ is the observations in the parent node and $T_1$ and $T_2$ the observations in the two subsets, $\bar{T}, \bar{T}_1$, and $\bar{T}_2$ are the arithmetical averages of the created final leaves. The risk of overfitting can be mitigated by introducing a minimum amount of observations required in every node or by introducing a second, shrinking algorithm, reducing the tree. In this study no shrinking algorithm has been tested but the number of observations in each leaf has been restricted to 7.5 % of the total sample which is used as a stopping criterion instead of a minimum level of standard deviation reduction.

F-TREE (FT)

An alternative way of creating a tree model, to the authors’ knowledge not proposed in academic literature, is to base the creation on OLS regression results or rather the significance of the independent variables in an OLS regression. When regressing an independent variable on a dependent one it is easy to calculate a standard error and from there an F-statistic for the independent variable. The F-statistic can then be used to generate a p-value for the hypothesis that the effect of the independent variable is zero. That is, that the independent variable does not affect the dependent variable. The F-tree is created by always splitting on the independent variable with the highest F-statistic (lowest p-value). The F-statistic is calculated from a regression with only one independent variable present in the model. Contrary to the regression tree, this model can only utilize dummy variables since it includes no algorithm for determining where to split a continuous variable. However, continuous variables can be included in the model if they are converted into dummy variables beforehand. In this study this has been done by testing different possible splits and then comparing the p-values of a linear regression on the dummy variable.

Although the F-tree only creates small leaves if they are significantly different, small leaves can be problematic from a practical viewpoint. The main problem with small leaves is that a few new observations may substantially change the estimated LGD for the leaf since the estimated value is just the arithmetical average.

HISTORICAL AVERAGE (HIST)

Instead of using a sophisticated model, one could simply use the historical average as a prediction of the future LGD levels. This method is included in the study as a benchmark.
4 DATA

As suggested by the Basel committee this study utilizes pooling of data as a remedy for the small amount of observations. The data was sourced from a member bank of PECDC which has access to a subset of the total database. The PECDC database is the world’s largest LGD/EAD database and members receive data in return for submitting data of the same type and year of default. PECDC members consist of 40 banks from Europe, Africa, North America, Asia and Australia (PECDC 2013b). While this study would have been impossible to conduct without the pooling of data such a procedure does have limitations including possible inconsistency in definitions and practice between organisations which reduce the comparability of data from different providers.

The subset of the database consists of around 60 000 observations ranging from 1983 to 2012 but with the vast majority occurring during the second half of this time span. However, restricting the data to financial institutions leaves a sample of less than 1 000 observations occurring between 1992 and 2012. It is necessary to do this restriction since the model is supposed to be suitable for exposures towards these kinds of counterparties and they are likely to differ from other companies. The low default portfolio problem gets even more severe since the data is incomplete. The low number of defaults for years prior to 2000 reflects the shortage of data during these years.

Due to the low number of observations it would have been beneficial to be able to include unfinished bankruptcy processes but unfortunately this data has not been available. Including only completed workout processes can potentially lead to a biased LGD estimation due to the underrepresentation of defaults with a longer process. The problem arises from the positive correlation between the length of the workout process and the LGD level and is likely to be more severe the shorter the sample time (Gürtler and Hibbeln, 2013). The problem is sometimes referred to as the resolution bias (PECDC, 2013a). In order to mitigate the resolution bias, the two most recent years (2011 and 2012) are excluded from the data. Two years has been deemed a reasonable time period by considering the time to resolution in the sample, see Table 3. In order to determine the time period to remove, the length of the resolution times for observations with a nominal LGD larger than zero has been analyzed. Observations with a nominal LGD equal to zero are not considered since some of these observations have a very short resolution time. The resolution bias might still be present but due to the overall shortage of data, there is a trade-off between removing recent data and thereby mitigating the resolution bias and keeping a sample big enough to base a model on.

<table>
<thead>
<tr>
<th>Proportion of observations with time to resolution shorter than</th>
<th>All obs.</th>
<th>Obs. with nominal LGD &gt; 0</th>
</tr>
</thead>
<tbody>
<tr>
<td>1 year</td>
<td>35%</td>
<td>25%</td>
</tr>
<tr>
<td>2 year</td>
<td>65%</td>
<td>55%</td>
</tr>
<tr>
<td>3 year</td>
<td>75%</td>
<td>65%</td>
</tr>
<tr>
<td>4 year</td>
<td>85%</td>
<td>80%</td>
</tr>
</tbody>
</table>

Table 3. Time to resolution (illustrative figures)

PECDC advises that the data is subject to validation filters as it is input by the banks and also to audits and reasonableness checks during the aggregation process. However, as with all data it could contain errors and therefore it has been searched for abnormal entries. In addition to the observations with a default date during 2011 and 2012, observations with a defaulted amount smaller than 1 EUR and observations with a collateral value more than 5 times more valuable than the defaulted amount have been excluded. Furthermore, a few observations with percentage guaranteed values below or equal to 1% has
been multiplied with 100 since they are believed to be typos. It seems unlikely that someone should guarantee only 1% or less of the amount.

In addition, observations with unknown facility type and facility types with less than 5 observations have been excluded. The model is based on observations of defaulted bonds, loans, revolvers, overdrafts, payment guarantees and derivative or securities claims and should be applied on exposures of these types only. For exposures of other types, a qualitative judgement regarding the similarity to these types of exposures must be conducted. Figure 2 shows the number of observations per year for both the initial data and for the remaining financial institutions only.

The remaining data bears the characteristics of a bimodal distribution with a higher mode for lower LGD rates that is often reported (Schuermann, 2004), see Figure 3. As mentioned earlier, the sample in this study consists of various exposures to other financial institutions. Since these companies are typically large and have a very good credit quality it can be classified as a so called low default portfolio. However, the data set used in this study lacks observations of e.g. defaulted covered bonds and repos which are often found in LDPs.

The geographical dispersion of the remaining sample is presented in Table 4.
<table>
<thead>
<tr>
<th>Country</th>
<th># obs.</th>
<th>Average LGD (%)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Germany</td>
<td>200</td>
<td>26.3</td>
</tr>
<tr>
<td>Denmark</td>
<td>110</td>
<td>18.5</td>
</tr>
<tr>
<td>Unknown</td>
<td>80</td>
<td>61.0</td>
</tr>
<tr>
<td>US</td>
<td>46</td>
<td>32.8</td>
</tr>
<tr>
<td>Kazakhstan</td>
<td>43</td>
<td>31.4</td>
</tr>
<tr>
<td>Iceland</td>
<td>30</td>
<td>65.2</td>
</tr>
<tr>
<td>UK</td>
<td>23</td>
<td>9.7</td>
</tr>
<tr>
<td>France</td>
<td>20</td>
<td>38.3</td>
</tr>
<tr>
<td>Ukraine</td>
<td>17</td>
<td>23.4</td>
</tr>
<tr>
<td>Turkmenistan</td>
<td>12</td>
<td>3.1</td>
</tr>
<tr>
<td>Norway</td>
<td>6</td>
<td>56.2</td>
</tr>
<tr>
<td>Argentina</td>
<td>6</td>
<td>37.1</td>
</tr>
<tr>
<td>Indonesia</td>
<td>4</td>
<td>19.4</td>
</tr>
<tr>
<td>Netherlands</td>
<td>4</td>
<td>46.4</td>
</tr>
<tr>
<td>Russia</td>
<td>4</td>
<td>50.5</td>
</tr>
<tr>
<td>Finland</td>
<td>0</td>
<td>-</td>
</tr>
<tr>
<td>Sweden</td>
<td>0</td>
<td>-</td>
</tr>
<tr>
<td>Other</td>
<td>71</td>
<td>38.9</td>
</tr>
<tr>
<td></td>
<td>676</td>
<td>32.8</td>
</tr>
</tbody>
</table>

Table 4. Country of jurisdiction for financial institutions (illustrative figures)
5 MODEL EVALUATION METHODS

WITHIN SAMPLE

The within sample testing evaluates the models’ power to predict LGD levels on the same sample as the parameters was based upon, i.e. the sample with financial institutions presented in Section 4. It can be looked upon as a measure of the models’ sample fit.

OUT OF SAMPLE

The out of sample testing follows the approach outlined by Bastos (2010), a so called \( k \)-fold cross-validation. Due to the small sample a \( 5 \)-fold cross validation is employed in this study instead of the \( 10 \)-fold cross validation used by Bastos (2010). The \( 5 \)-fold cross validation splits the sample into five roughly equal parts and the parameters of the model are estimated based on four of these five subsets. The ME, MAE and R2 values are then calculated based on predictions on the remaining part. The procedure is then repeated for the four other subsets and average ME, MAE and R2 values are calculated. The whole procedure, including splitting the sample, is iterated 100 times in order to receive a more stable estimate.

For the two tree models, the tree structure is treated as constant and the structure estimated from the whole sample is used also in the out of sample testing.

OUT OF TIME

Out of time testing, or back testing, is a common technique for evaluating risk models. In a LDP setting, it has however severe limitations due to the lack of extensive data. ISDA (2005) notes that “for the majority of LDP models, the results of a back testing exercise will not provide any appropriate evidence to support the IRB approval process”. Despite the limitations of out of time testing on a LDP portfolio an out of time evaluation has been conducted. The sample has been into three periods, prior to 2008, 2008 and 2009-2010. The first subset has then been used to estimate a model which predictive power has been tested on the second subset. The first and second subset has then been used to estimate a model which predictive power has been tested on the third subset.

Similar to the out of sample testing, the tree structure is treated as constant and the same structure as derived in the within sample testing is used also in the out of time testing.

<table>
<thead>
<tr>
<th>Subset</th>
<th>Years</th>
<th># obs.</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>1991-2008</td>
<td>400</td>
</tr>
<tr>
<td>2</td>
<td>2008</td>
<td>120</td>
</tr>
<tr>
<td>3</td>
<td>2009-2010</td>
<td>156</td>
</tr>
</tbody>
</table>

<p>| | | |</p>
<table>
<thead>
<tr>
<th></th>
<th></th>
<th></th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td>676</td>
</tr>
</tbody>
</table>

Table 5. Subsets for out of time testing (illustrative figures)
MEASURES OF PREDICTIVE POWER

To evaluate the LGD estimation techniques three measures have been used. For each technique the performance is measured in mean error (ME), mean absolute error (MAE) and R-squared value (R2). The ME is expressed as

$$ME = \frac{1}{n} \sum_{i=1}^{n} \hat{\text{LGD}}_i - \text{LGD}_i$$

where n is the number of observations, $\hat{\text{LGD}}_i$ is the estimated value of LGD for the exposure i with a given model and LGD$_i$ is the observed LGD value for exposure i. While the average error gives an indication of whether the model is biased, MAE shows the size of the errors. MAE is defined as

$$MAE = \frac{1}{n} \sum_{i=1}^{n} |\hat{\text{LGD}}_i - \text{LGD}_i|,$$

with definitions as above. Finally R2 is defined as

$$R^2 = 1 - \frac{\sum_{i=1}^{n} (\hat{\text{LGD}}_i - \text{LGD}_i)^2}{\sum_{i=1}^{n} (\hat{\text{LGD}} - \text{LGD})^2},$$

where $\text{LGD}$ is the average LGD in the sample. Hence R2 measures the percentage of variations that can be explained by the model. While R2 is bounded between zero and one within sample R2 can become negative if the model is actually worse than using the average in out of sample and out of time testing (Loterman et al., 2012). To give a measure of how stable the ME, MAE and R2 values are the standard deviations of the measures are also calculated and displayed in brackets after the values.

LGD DISTRIBUTIONS FOR EXAMPLE PORTFOLIO

Finally the models are tested by bootstrapping data and evaluating the LGD levels on an example portfolio consisting of one exposure of every possible combination, in total 80 different exposures. The bootstrapping is repeated 1 000 times to produce a distribution of the LGD estimates for each of the tested models. This bootstrapping is also used to calculate the mean, 1st and 99th percentile of the resulting LGD distribution for the example portfolio.

CORRELATION MATRIX OF LGD ESTIMATES

In order to given an indication of the similarity of the LGD estimates for different exposures resulting from the different models a correlation matrix is constructed. The correlation matrix is based on the LGD estimates for the exposures in the example portfolio defined in the section above. Since the 80 exposures in the example portfolio receives a LGD estimate for each of the six models the correlations between the LGD estimates are calculated and displayed in a matrix.
6 Results

6.1 Risk Drivers

The following risk drivers have been considered as explanatory variables. The number of risk drivers in a credit risk model can vary and sometimes include up to 30-40 inputs when the amount of publically available data is large (ISDA, 2005). Since the observations in the database are anonymous, it has not been possible to enrich the data with additional borrower information and the number of risk drivers is rather low. Furthermore, risk drivers without support in the data and in theory have been left out of the final models. This analysis has been conducted using OLS regressions. Continuous risk drivers have been bounded to the range 0 to 1 in order to make the contribution to the estimation more clear. The risk drivers included in the final models are collateral, guarantee, industry, geographic region, default year as well as the facility types overdraft, revolver, payment guarantee and loan.

Macroeconomic Environment

Several macroeconomic variables have been tested in the OLS model, see Table 6. Domestic macroeconomic variables are not considered since it is believed that the state of the global financial market is more important than the state of the domestic economy for exposures towards financial institutions. The basic idea of incorporating a macroeconomic variable is that it should capture the effect of higher LGD levels during bad economic times reported in several studies. That would mean a negative effect on the LGD from the macroeconomic variable. While this makes sense from a theoretical viewpoint the reverse relationship is counterintuitive and difficult to justify. However, all macroeconomic variables receive a positive parameter in the OLS model. One explanation for these results is believed to be the downward sloping trend in both the LGD levels and many of the macroeconomic variables. Also other studies have also experienced problems in capturing the believed relationship between LGD levels and the macroeconomic environment. A study by PECD (2013a) did find a negative relationship between LGD levels and the OECD GDP growth rate but only when the specific timing of the recovery cash flows was taken into account. Since the timing of these cash flows as well as the future GDP growth is unknown at, and prior to, the event of default it cannot be included in a model intended for practical usage.
Estimation of Loss Given Default for Low Default Portfolios

With separate linear trend | Without separate linear trend
---|---
**Macroeconomic variables** | **Effect** | **p-value** | **Effect** | **p-value**
Germany stock market return (1Y) | - | >5% | + | >5%
UK stock market return (1Y) | - | <5% | + | >5%
US stock market return (1Y) | - | >5% | + | >5%
Euro area stock market return (1Y) | - | >5% | + | >5%
Short term interest rate UK | + | <5% | + | <5%
Short term interest rate US | + | >5% | + | <5%
Short term interest rate Germany | + | <5% | + | <5%
Long term interest rate UK | + | <5% | + | <5%
Long term interest rate US | + | >5% | + | <5%
Long term interest rate Euro area | + | >5% | + | <5%
Long term interest rate Germany | + | <5% | + | <5%
Long term interest rate Germany (trend removed) | - | >5%
GDP growth OECD | + | >5% | + | <5%
GDP growth US | + | >5% | + | <5%
GDP growth Germany | + | 1 | + | 0

Table 6. Macroeconomic variables in OLS regression

The effect of the downward sloping trend in the macroeconomic variables and the LGD levels can be illustrated with the example of the German long term interest rate. The interest rate has a positive effect on the LGD levels, and the p-value is below 5%, but if the trend component is removed, see Figure 4, the effect becomes negative instead. The trend is removed monthly, based on the regression parameter from a regression of the interest rate on a straight line.

![Figure 4. Trend component in German long term interest rate (Data source: OECD)](image)

The effect of the downward sloping trend in the macroeconomic variables and the LGD levels can be illustrated with the example of the German long term interest rate. The interest rate has a positive effect on the LGD levels, and the p-value is below 5%, but if the trend component is removed, see Figure 4, the effect becomes negative instead. The trend is removed monthly, based on the regression parameter from a regression of the interest rate on a straight line.

![Figure 5. UK 1Y stock market returns (Data source: OECD)](image)
Instead of removing the trend from the macroeconomic time series one could incorporate a separate linear trend in the regression in order to capture the decreasing trend in the LGD levels. Such a trend has a significant negative effect on the estimated LGD levels in an OLS regression. If such a trend is incorporated the stock market return parameters change sign and have a negative effect on the LGD levels. The other variables still have a positive effect but the p-values increase and many are now not significant, see Table 6. The most significant stock market return parameter is the UK stock market return. This time series, showed in Figure 5, bears more characteristics of a business cycle proxy than for example the long term interest rates in Figure 4. The stock market returns and the interest rates have the advantage of being publically available at every point in time unlike GDP growth which is only known subsequently. It could also be argued that both stock market return and interest rate figures are forward looking to some extent which would be a benefit since it is the macroeconomic environment during the bankruptcy process and not at the default date that is believed to influence the LGD levels.

However, in a practical model, a macroeconomic variable sometimes results in unintuitive splits in the tree models with higher LGDs during supposed better economic times. Secondly, tree models sometimes create leaves based on very small differences in the macroeconomic variables which seem unlikely to hold out of sample. Furthermore, as previously mentioned the LGD used in the report of credit risk must be a so called “Downturn LGD”, i.e. reflecting the LGD in a downturn environment. Because of this the effect of the macroeconomic variable needs to be both large enough in magnitude to create a substantial difference during economic downturns and affect observations of all kinds. The macroeconomic variable introduced here fails to have a substantial effect in an OLS model and does only affect parts of the tree models since the splits are too far down in the trees. Due to these reasons no macroeconomic variable was used in the final models.

**SENIORITY & FACILITY TYPE**

Several studies have found the seniority of the claim to be one of the most important determinants of LGD. However, the seniority parameter failed to prove significant in an OLS regression and actually indicated a positive relationship between LGD and seniority. It was therefore dropped from the models.

Most academic LGD studies focus on loans and bonds but since also other facility types differ in usage and risk profile, the facility type can be used as a risk driver. In this sample, the dummy variables for the facility types payment guarantee, overdraft, and revolver has been found to significantly affect the LGD levels in an OLS regression. However, the dummy variables for the types bond, loan and derivatives and securities claim failed to prove significant in the sample. In the F-tree model the loan dummy is once again included since it has been proved significant in subsets of the data.

**COLLATERAL**

According to the academic literature collateral should be an important determiner of the realized LGD. While this effect was found in the sample the effect was not as significant as for other variables and was sensitive to specifications of the variable. It is not only the size of the collateral which matters for the realized recovery. While financial collateral usually can be converted to cash easily, physical collateral can be cumbersome to sell, especially to a fair value, since the number of buyers can be few and the market illiquid. To mitigate this problem the variable used in the models is the percentage of financial and cash collateral out of the defaulted amount. Other types of collaterals are not used in the models. The variable is also capped at 100 %, meaning that any collateral worth more than 100 % of the defaulted amount will still only count as 100%.
**GEOGRAPHIC REGION (DEVELOPED COUNTRY)**

A model utilizing exposures from all over the world needs to check for systematic differences between geographic regions. Several ways of grouping countries were tested (including EU, EEA, Euro zone, North America, Scandinavia, OECD, Emerging markets etc.). The distinction between developed and developing countries was chosen for several reasons. First, it yields two substantial groups in the sample. Secondly, it is intuitive and there is no need to make judgemental decisions regarding for instance whether Denmark should be grouped with the Euro countries or not. It also gave reasonable and significant results in an OLS regression and no residual group with countries not belonging to any group appeared. The variable in the models is a dummy variable indicating whether the borrower’s country of jurisdiction is a developed country or not. If the country of jurisdiction is unknown the country of residence is used instead.

**INDUSTRY (BANKS)**

Since this study considers only the financial industry grouping of observations on industry level is not possible. However some differences regarding the type of borrower can be seen in the data. Since banks face higher regulatory requirements than other financial institutions it could be supposed that this would influence the realized LGD levels. Another explanation could be that banks are more likely to be saved when facing default due to their importance to the economy. The data gives some support to this theory when looking at the proportion of nominal LGDs equalling zero. The variable used in the models is a dummy variable indicating whether the counterparty is a bank or not.

<table>
<thead>
<tr>
<th>Banks</th>
<th>Non-bank Financial institutions</th>
</tr>
</thead>
<tbody>
<tr>
<td>Nominal LGDs = 0</td>
<td>60%</td>
</tr>
</tbody>
</table>

Table 7. Bank and non-banks with nominal LGD equal to zero (Illustrative figures)

**SIZE OF FIRM & EXPOSURE**

Many academic studies use the size of the exposure as an explanatory variable for LGD levels (Bastos, 2010; Khieu et al., 2012). While this could help explaining LGD for e.g. SMEs, where a company has just one or a few lenders, it is probably not useful for the case of a low default portfolio with a lot of banks and institutions since these entities typically have many liabilities to a huge amount of counterparties. The argument that higher default amounts lead to lower recovery rates as banks are unwilling to push larger loans to default resulting in lower recoveries when loans actually default which has sometimes been proposed (Khieu et al. 2012) seems unlikely to be valid for the kind of obligors in this data. Since banks and other financial institutions have a huge amount of creditors and exposures it is usually not up to one single creditor whether to push the institution to default. Furthermore, including the size of the exposure is also problematic from a practical point of view. It is hard to justify for business units why two small loans should have a larger (or smaller) expected loss than one big. The size of the firm has not been tested as an explanatory variable for realized LGD levels since it is in most cases not reported to the database in order to ensure the anonymity of the data.

**INDUSTRY DISTRESS**

Industry distress as a risk driver is not considered since only the financial industry is included in this study and it is believed that the state of the economy can serve as a reasonable proxy for the state of the financial industry. The use of macroeconomic variables as a proxy for the degree of distress in the financial industry can be motivated by the fact that
financial crises often lead to severe recessions (Reinhart and Rogoff, 2009). The findings by Cebula et al (2011) indicating a positive relationship between the growth rate of real GDP and the failure rate of banks further supports the use of this proxy. It also seems intuitive from a theoretical viewpoint since a stronger economy should result in a stronger performance of bank loans, reducing the risk of bank failures.

**LEVERAGE**

Similar to firm size, the borrower’s leverage is not included in the database.

**GUARANTEE**

A guarantee from a third party is expected to decrease the LGD levels. This effect has been proved significant in an OLS regression. The variable in the models is the percentage of the defaulted amount which is guaranteed. The variable is capped at 100% meaning that observations with a higher percentage guaranteed still receive the value 100%.

**UTILISATION RATE**

A variable capturing the utilization rate at the event of default failed to prove significant in an OLS regression. Since it is also likely that the utilisation rate soars just before the event of default it is not appropriate to use as a risk driver from a practical viewpoint. Utilisation rate was therefore dropped from the models.
Default Year (Pre-2005 Dummy)

Incorporating a yearly trend in the data can be problematic from a practical point of view since the model would then suggest that the LGD levels are decreasing in time. If the model is to be used in practice all predictions of future LGDs would be adjusted downwards with some constant times the number of years in the sample, 20 years in this case. This is problematic for two reasons. First, in this study the adjustment is rather large in magnitude and may thereby be hard to justify. Secondly the model would run into problems if the LGD levels start to increase again and the linear effect disappears. Instead of using a linear trend throughout the sample period a dummy variable for early observations can be included. This approach has several advantages. Firstly it is easier to justify since the effect is restricted to early observations and not the whole sample. Secondly the effect will fade away when more observations are included and no problems will arise if the linear downward sloping trend in the data disappears with new observations. Thirdly, a pre-2005 dummy can be justified from a theoretical viewpoint since this was the year when the Basel default definitions were released. One could suspect that before these definitions were effective, the default definitions were stricter and too few observations with a zero loss were reported as defaults. This is however not supported by the data in this study. The number of observations where the nominal LGD equals zero does not seem to increase over time, see Figure 6. However, the choice of 2005 for the dummy variable still seems reasonable when looking at the data. Table 8 shows p-values from the OLS model for different choices of year. The year 2005 gives a very low p-value for the dummy variable and there are still a substantial number of observations after this year.

<table>
<thead>
<tr>
<th>Year</th>
<th>p-value</th>
<th># obs. later than</th>
</tr>
</thead>
<tbody>
<tr>
<td>1991</td>
<td>&gt;10%</td>
<td>676</td>
</tr>
<tr>
<td>1992</td>
<td>&gt;10%</td>
<td>674</td>
</tr>
<tr>
<td>1993</td>
<td>&lt;1%</td>
<td>670</td>
</tr>
<tr>
<td>1995</td>
<td>&lt;1%</td>
<td>660</td>
</tr>
<tr>
<td>1996</td>
<td>&lt;1%</td>
<td>657</td>
</tr>
<tr>
<td>1997</td>
<td>&lt;1%</td>
<td>655</td>
</tr>
<tr>
<td>1998</td>
<td>&lt;1%</td>
<td>652</td>
</tr>
<tr>
<td>1999</td>
<td>&lt;10%</td>
<td>530</td>
</tr>
<tr>
<td>2000</td>
<td>&lt;10%</td>
<td>504</td>
</tr>
<tr>
<td>2001</td>
<td>&lt;10%</td>
<td>485</td>
</tr>
<tr>
<td>2002</td>
<td>&lt;1%</td>
<td>460</td>
</tr>
<tr>
<td>2003</td>
<td>&lt;1%</td>
<td>420</td>
</tr>
<tr>
<td>2004</td>
<td>&lt;1%</td>
<td>392</td>
</tr>
<tr>
<td>2005</td>
<td>&lt;1%</td>
<td>365</td>
</tr>
<tr>
<td>2006</td>
<td>&lt;1%</td>
<td>340</td>
</tr>
<tr>
<td>2007</td>
<td>&lt;1%</td>
<td>310</td>
</tr>
<tr>
<td>2008</td>
<td>&lt;1%</td>
<td>276</td>
</tr>
<tr>
<td>2009</td>
<td>&gt;10%</td>
<td>156</td>
</tr>
</tbody>
</table>

Table 8. Splitting year into dummy variable (illustrative figure)

![Figure 6. Proportion of nominal LGD equal to zero](image)

While it seems reasonable from a mathematical point of view to try to include all observations and capture differences in reporting habits it can be problematic to justify that the average LGD level in the sample is too high and should be adjusted downwards. The test results are therefore reported both with and without such an adjustment.
6.2 Final Models

All models are reported both without the Pre 2005 dummy (Original models) and with the Pre 2005 dummy (Adjusted models).

**Ordinary Least Squares Regression (OLS)**

The model with original LGD observations gives negative parameters for the variables Payment Guarantee, Revolver, Financial/Cash Collateral, Developed Country, Bank, Guarantee and a positive parameter for the Overdraft variable. All parameters except for the Collateral variable are significant on the 1% level, while the Collateral parameter is significant on the 5% level.

The model with adjusted LGD observations gives similar parameters and significance levels as when original LGD observations are used. However, the Overdraft parameter is only significant at the 10% level in this case. The Pre 2005 dummy receives a significant positive parameter.

<table>
<thead>
<tr>
<th>OLS Regression</th>
<th>Pre 2005 dummy</th>
<th>Payment guarantee</th>
<th>Overdraft</th>
<th>Revolver</th>
<th>% Financial /Cash collateral</th>
<th>Developed Country</th>
<th>Bank</th>
<th>Guarantee (%)</th>
<th>Intercept</th>
</tr>
</thead>
<tbody>
<tr>
<td>Original OLS parameters</td>
<td>-17.2</td>
<td>19.3</td>
<td>-17.3</td>
<td>-11.0</td>
<td>-19.1</td>
<td>-14.8</td>
<td>-21.4</td>
<td>84.3</td>
<td></td>
</tr>
<tr>
<td>Significance level Original</td>
<td>&lt;1%</td>
<td>&lt;1%</td>
<td>&lt;1%</td>
<td>&lt;5%</td>
<td>&lt;1%</td>
<td>&lt;1%</td>
<td>&lt;1%</td>
<td>&lt;1%</td>
<td></td>
</tr>
<tr>
<td>Adjusted OLS parameter</td>
<td>16.3</td>
<td>-19.1</td>
<td>14.6</td>
<td>-21.2</td>
<td>-10.8</td>
<td>-19.7</td>
<td>-14.2</td>
<td>-22.3</td>
<td>80.1</td>
</tr>
<tr>
<td>Significance level Adjusted</td>
<td>&lt;1%</td>
<td>&lt;1%</td>
<td>&lt;10%</td>
<td>&lt;1%</td>
<td>&lt;5%</td>
<td>&lt;1%</td>
<td>&lt;1%</td>
<td>&lt;1%</td>
<td>&lt;1%</td>
</tr>
</tbody>
</table>

Table 9. OLS regression results (illustrative figures)

**Ridge Regression (RiR)**

All regression parameters in both the adjusted and the original LGD models receive the same signs and are similar in magnitude to the OLS model.

<table>
<thead>
<tr>
<th>Ridge Regression</th>
<th>Pre 2005 dummy</th>
<th>Payment guarantee</th>
<th>Overdraft</th>
<th>Revolver</th>
<th>% Financial /Cash collateral</th>
<th>Developed Country</th>
<th>Bank</th>
<th>Guarantee (%)</th>
<th>Intercept</th>
</tr>
</thead>
<tbody>
<tr>
<td>Original RiR Parameters</td>
<td>-16.2</td>
<td>19.3</td>
<td>-16.3</td>
<td>-10.2</td>
<td>-18.2</td>
<td>-12.1</td>
<td>-20.3</td>
<td>79.3</td>
<td></td>
</tr>
<tr>
<td>Adjusted RiR Parameters</td>
<td>16.2</td>
<td>-18.2</td>
<td>14.2</td>
<td>-20.1</td>
<td>-19.2</td>
<td>-20.7</td>
<td>-11.2</td>
<td>-21.3</td>
<td>78.1</td>
</tr>
</tbody>
</table>

Table 10. Ridge regression results (illustrative figures)

**Logistic-OLS Regressions Model (LR-OLS)**

In the Logistic-OLS Regressions model with original LGD observations the probability of LGD=0 is increased by all variables except the Overdraft variable. For the probability of LGD=1 the signs are reversed for all variables except Payment Guarantee. That is, only the Overdraft and Payment Guarantee variables increase the probability for LGD=1. The OLS regression for values between 0 and 1 yields a positive parameter for the Overdraft variable and negative parameters for all other variables.

In the Logistic-OLS Regressions model with adjusted LGD observations, the probability of LGD=0 is increased by the variables Payment Guarantee, Revolver, Collateral, Developed Country, Bank Guarantee and the Pre-2005 dummy while the Overdraft variable decreases the probability. For the probability of LGD=1 only the Overdraft, Payment Guarantee and Pre 2005 dummy variable
increase the probability. The Pre 2005 dummy was expected to receive a negative parameter in the regression for P(LGD=0) as more observations of no loss is expected to be included in the data set after 2005. This might be an indication that the theoretical reason for including the Pre 2005 dummy is not supported by the data. The OLS regression shows a positive parameter only for the Pre 2005 dummy variable, all other parameters have negative signs.

<table>
<thead>
<tr>
<th>Logistic-OLS Regressions model</th>
<th>Pre 2005 dummy</th>
<th>Payment guarantee</th>
<th>Overdraft</th>
<th>Revolver</th>
<th>% Financial /Cash collateral</th>
<th>Developed Country</th>
<th>Bank</th>
<th>Guarantee (%)</th>
<th>Intercept</th>
</tr>
</thead>
<tbody>
<tr>
<td>Original 0 &lt; LGD &lt;1</td>
<td>-8.3</td>
<td>8.5</td>
<td>-9.2</td>
<td>-10.2</td>
<td>-4.6</td>
<td>-9.2</td>
<td>-15.5</td>
<td>61.5</td>
<td></td>
</tr>
<tr>
<td>Original P(LGD=0)</td>
<td>3.5</td>
<td>-0.8</td>
<td>2.4</td>
<td>0.5</td>
<td>3.1</td>
<td>2.2</td>
<td>2.6</td>
<td>-7.4</td>
<td></td>
</tr>
<tr>
<td>Original P(LGD=1)</td>
<td>0.1</td>
<td>1.4</td>
<td>-0.7</td>
<td>-0.7</td>
<td>-1.1</td>
<td>-0.8</td>
<td>-2.3</td>
<td>-0.1</td>
<td></td>
</tr>
<tr>
<td>Adjusted 0 &lt; LGD &lt;1</td>
<td>5.3</td>
<td>-7.3</td>
<td>-0.3</td>
<td>-10.3</td>
<td>-11.1</td>
<td>-2.7</td>
<td>-7.0</td>
<td>-15.1</td>
<td></td>
</tr>
<tr>
<td>Adjusted P(LGD=0)</td>
<td>1.7</td>
<td>4.0</td>
<td>-2.3</td>
<td>2.3</td>
<td>0.5</td>
<td>2.8</td>
<td>1.6</td>
<td>2.5</td>
<td></td>
</tr>
<tr>
<td>Adjusted P(LGD=1)</td>
<td>2.3</td>
<td>0.7</td>
<td>0.5</td>
<td>-1.0</td>
<td>-1.5</td>
<td>-1.3</td>
<td>-0.7</td>
<td>-2.3</td>
<td></td>
</tr>
</tbody>
</table>

Table 11. Logistic-OLS Regressions results (illustrative figures)

**TRIMMED LOGISTIC-OLS REGRESSIONS (TLR-OLS)**

In the Trimmed Logistic-OLS Regressions model the probability of LGD=0 is equal to the probability of LGD=0 in the LR-OLS model above, that is the regressions are identical. The OLS regression for the original observations gives a positive parameter for the Overdraft variable and negative parameters for the other variables, just like in the LR-OLS model. In the adjusted model the Pre 2005 dummy has a positive parameter in the OLS regression. All other effects are in line with the original model but the positive effect of Overdraft is much lower in the adjusted LGD model. The parameters for all variables apart from the Overdraft variable receive the same signs in the TLR-OLS model as in the LR-OLS model but most differ in magnitude.

<table>
<thead>
<tr>
<th>Trimmed Logistic-OLS Regressions model</th>
<th>Pre 2005 dummy</th>
<th>Payment guarantee</th>
<th>Overdraft</th>
<th>Revolver</th>
<th>% Financial /Cash collateral</th>
<th>Developed Country</th>
<th>Bank</th>
<th>Guarantee (%)</th>
<th>Intercept</th>
</tr>
</thead>
<tbody>
<tr>
<td>Original 0 &lt; LGD</td>
<td>-5.5</td>
<td>20.7</td>
<td>-13.8</td>
<td>-14.0</td>
<td>-12.9</td>
<td>-13.6</td>
<td>-25.2</td>
<td>79.1</td>
<td></td>
</tr>
<tr>
<td>Original P (LGD=0)</td>
<td>3.5</td>
<td>-0.8</td>
<td>2.4</td>
<td>0.5</td>
<td>3.1</td>
<td>2.2</td>
<td>2.6</td>
<td>-7.4</td>
<td></td>
</tr>
<tr>
<td>Adjusted 0 &lt; LGD</td>
<td>18.3</td>
<td>-1.2</td>
<td>7.2</td>
<td>-15.2</td>
<td>-15.3</td>
<td>-10.2</td>
<td>-26.3</td>
<td>64.4</td>
<td></td>
</tr>
<tr>
<td>Adjusted P (LGD=0)</td>
<td>1.7</td>
<td>4.0</td>
<td>-2.3</td>
<td>2.3</td>
<td>0.5</td>
<td>2.8</td>
<td>1.6</td>
<td>2.5</td>
<td></td>
</tr>
</tbody>
</table>

Table 12. Trimmed Logistic-OLS Regressions results (illustrative figures)

**REGRESSION TREE (RT)**

The regression tree generated with original LGD observations, Figure 7, starts with splits on the product types Overdraft and Revolver. If the observation is a revolver the sample is then split upon Developed/Developing Country. If the observation is neither an overdraft nor a revolver the sample is split upon the level of guarantee. The observations with low guarantee levels are then split in Payment Guarantees and Non Payment Guarantees.

The regression tree with adjusted LGD observations, Figure 8, starts with the same 4 splits. The observations in the group with low guarantees are however split in Developed and Developing Countries before splitting on Payment Guarantees.
Estimation of Loss Given Default for Low Default Portfolios

F-TREE (FT)

The F-tree in Figure 9 is based on the original LGD values, that is, early observations have not been adjusted. The data is split into 9 different leaves. The F-tree in Figure 10 has been created based on adjusted LGD values prior to 2005. Here the data is split into 10 different leaves. Both models include the Loan product type which failed to prove significant in the regression models. In general the two models are similar apart from an extra split on the guarantee level in the adjusted LGD model.

Both the F-tree and the Regression tree could easily be modified due to qualitative considerations without a large worsening in the test results. It could also be noted that the F-tree model produces larger trees than the RT model where the minimum leaf size is set to 7.5% of the sample.
6.3 Test results

The models have been evaluated based on the statistical techniques outlined in Section 5.

Predictive power

The results of the within sample, out of sample and out of time tests are presented in Table 13 and 14. Within sample the LR-OLS, TLR-OLS, FT and RT models give unbiased estimates while the RiR and OLS models have small positive biases due to the truncation of estimates to the [0,1] interval. This is true for the estimations based both on original and adjusted LGD observations for all models except LR-OLS which produces a small negative bias using adjusted LGD levels. In terms of MAE and R2 the FT model produces the best results within sample. However the differences between the models are not large. All models produce positive R2 values and substantially lower MAE than using historical average in the within sample test.

The errors increase out of sample but not to a large extent. In the out of sample test the FT model gives the lowest MAE regardless of whether observations are adjusted or not. In terms of ME all models seem to be close to unbiased. The original observations result in similar R2 values for all models except for RT which has a substantially lower R2 value. Using adjusted observations the LR-OLS and TLR-OLS models give higher R2 values than the other models. All models give lower errors and higher R2 values than using the Historical average approach out of sample.

The F-tree model gives the largest errors out of time, even larger than the historical average approach, both when observations are adjusted and not. The lowest errors are produced by the LR-OLS and TLR-OLS models when observations prior to 2005 are adjusted. The LR-OLS model is the only model to produce positive R2 values with original observations. In general no model produces substantially better results than using the Historical average approach in the out of time test. Once again, the limitations of the out of time test described in Section 5 needs to be taken into account.

Overall, errors are lower when observations prior to 2005 are adjusted.

<table>
<thead>
<tr>
<th></th>
<th>Within sample</th>
<th>Out of sample</th>
<th>Out of time</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>ME MAE R2</td>
<td>ME MAE R2</td>
<td>ME MAE R2</td>
</tr>
<tr>
<td>OLS</td>
<td>0.5 22.1 0.34</td>
<td>0.3 (4.3) 22.8 (1.6) 0.31 (0.14) 8.2 (11.8) 26.1 (2.8)</td>
<td>-0.16 (0.32)</td>
</tr>
<tr>
<td>RiR</td>
<td>0.3 22.1 0.34</td>
<td>0.2 (5.1) 22.8 (1.5) 0.31 (0.09) 7.8 (11.0) 26.1 (3.3)</td>
<td>-0.14 (0.30)</td>
</tr>
<tr>
<td>LR-OLS</td>
<td>0.0 21.4 0.36</td>
<td>0.1 (3.8) 21.9 (2.0) 0.32 (0.07) 8.3 (9.5) 26.0 (2.8)</td>
<td>0.13 (0.16)</td>
</tr>
<tr>
<td>TLR-OLS</td>
<td>0.0 21.3 0.36</td>
<td>0.0 (2.8) 21.7 (2.6) 0.32 (0.07) 7.8 (9.0) 26.0 (3.0)</td>
<td>-0.17 (0.32)</td>
</tr>
<tr>
<td>RT</td>
<td>0.0 24.7 0.23</td>
<td>0.2 (2.7) 24.8 (1.5) 0.24 (0.11) 2.8 (3.2) 25.4 (3.7)</td>
<td>-0.11 (0.29)</td>
</tr>
<tr>
<td>FT</td>
<td>0.0 20.9 0.36</td>
<td>-0.1 (3.2) 21.4 (2.0) 0.32 (0.09) 17.0 (11.5) 30.9 (0.2)</td>
<td>-0.61 (0.73)</td>
</tr>
<tr>
<td>Hist</td>
<td>0.0 28.9 0.00</td>
<td>-0.2 (4.7) 29.0 (1.3) -0.02 (0.04) 1.9 (5.3) 27.4 (6.0)</td>
<td>-0.05 (0.08)</td>
</tr>
</tbody>
</table>

Standard deviations in brackets

Table 13. Predictive power for original LGD levels (illustrative figures)
The estimated LGD levels depend of course on the model chosen. Table 16 summarizes LGD estimates for the example portfolio defined in Section 5 as well as for three specific exposures. The exposures are described in Table 17. The choice of model clearly influences the estimated LGD levels, especially for specific exposures. On a portfolio with different exposures the effects net out to some degree and the overall LGD estimate does not differ as much between different models. Since the tree models group exposures and assign the same LGD to the whole group the reasonability of the grouping needs to be justified from a theoretical viewpoint. The effect of the grouping on a real portfolio needs to be considered since many exposures might end up in the same leaf. This is especially important since the LGD estimates differ a lot in magnitude as seen in Table 16. Furthermore, tree models might fail to distinguish between exposures of rather different nature which could be problematic in practice.

Table 16. **LGD estimates (illustrative figures)**

<table>
<thead>
<tr>
<th>Example Portfolio</th>
<th>Exposure 1</th>
<th>Exposure 2</th>
<th>Exposure 3</th>
</tr>
</thead>
<tbody>
<tr>
<td>OLS</td>
<td>36.6</td>
<td>67.3</td>
<td>36.9</td>
</tr>
<tr>
<td>RiR</td>
<td>35.7</td>
<td>64.7</td>
<td>35.4</td>
</tr>
<tr>
<td>LR-OLS</td>
<td>31.6</td>
<td>55.9</td>
<td>36.9</td>
</tr>
<tr>
<td>TLR-OLS</td>
<td>33.4</td>
<td>61.1</td>
<td>38.8</td>
</tr>
<tr>
<td>RT</td>
<td>35.4</td>
<td>45.7</td>
<td>18.3</td>
</tr>
<tr>
<td>FT</td>
<td>33.8</td>
<td>70.5</td>
<td>17.0</td>
</tr>
<tr>
<td>Hist</td>
<td>32.9</td>
<td>32.9</td>
<td>32.9</td>
</tr>
</tbody>
</table>

Table 17. **Description of exposures in Table 16**

<table>
<thead>
<tr>
<th>Exposure</th>
<th>Description</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>Unsecured bond exposure towards bank in developed country</td>
</tr>
<tr>
<td>2</td>
<td>Collaterized loan to non-bank financial institution in developing country</td>
</tr>
<tr>
<td>3</td>
<td>Revolver to non-bank financial institution in developed country</td>
</tr>
</tbody>
</table>
LGD DISTRIBUTIONS FOR EXAMPLE PORTFOLIO

Bootstrapping the original data and recalculating the results for the example portfolio defined in Section 5 yields the distributions of LGD estimates presented in Table 18 and the percentiles presented in Table 19. The models using linear regressions generally results in a larger dispersion than the tree models.

<table>
<thead>
<tr>
<th>OLS</th>
<th>RiR</th>
<th>LR-OLS</th>
<th>TLR-OLS</th>
<th>RT</th>
<th>FT</th>
</tr>
</thead>
<tbody>
<tr>
<td>1st Percentile</td>
<td>29.3</td>
<td>28.6</td>
<td>24.4</td>
<td>26.8</td>
<td>28.1</td>
</tr>
<tr>
<td>Mean</td>
<td>36.8</td>
<td>35.9</td>
<td>31.4</td>
<td>35.6</td>
<td>31.8</td>
</tr>
<tr>
<td>99th Percentile</td>
<td>45.0</td>
<td>43.8</td>
<td>39.9</td>
<td>46.1</td>
<td>36.3</td>
</tr>
</tbody>
</table>

Table 18. LGD distribution for example portfolio (illustrative graphs)

<table>
<thead>
<tr>
<th>OLS</th>
<th>RiR</th>
<th>LR-OLS</th>
<th>TLR-OLS</th>
<th>RT</th>
<th>FT</th>
</tr>
</thead>
<tbody>
<tr>
<td>OLS</td>
<td>1.00</td>
<td>1.00</td>
<td>0.95</td>
<td>0.96</td>
<td>0.60</td>
</tr>
<tr>
<td>RiR</td>
<td>1.00</td>
<td>0.95</td>
<td>0.96</td>
<td>0.62</td>
<td>0.65</td>
</tr>
<tr>
<td>LR-OLS</td>
<td>1.00</td>
<td>0.99</td>
<td>0.56</td>
<td>0.71</td>
<td></td>
</tr>
<tr>
<td>TLR-OLS</td>
<td>1.00</td>
<td>0.60</td>
<td>0.71</td>
<td></td>
<td></td>
</tr>
<tr>
<td>RT</td>
<td>1.00</td>
<td>0.51</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>FT</td>
<td>1.00</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Table 20. Correlation matrix of LGD estimates in example portfolio (illustrative figures)

CORRELATION MATRIX OF LGD ESTIMATES

A correlation matrix of the LGD estimates produced by the example portfolio is presented in Table 20. As expected there is a strong positive correlation between the models. The correlation is however stronger between the models based on linear regression. The correlation between the two tree models is the lowest of all correlations in the matrix, while all correlations between based on linear regression are close to 1.

USING MODELS IN PRACTICE

A major difference between the tree models, RT and FT models, and the models based on linear regressions (OLS, RiR, LR-OLS and TLR-OLS) is that linear regressions give rise to an infinite set of possible LGD estimates. As a portfolio might consist of exposures with characteristics not found in the sample it is difficult to get an overview of the possible outcomes of LGD estimates. In the tree models all possible outcomes of LGD estimates are stated beforehand. However, consistency in ranking is not assured in the tree models. It is possible that “better” exposures receive a higher LGD estimate since they happen to be grouped with “worse” exposures having high LGDs. In the models including linear regressions (OLS, RiR, LR-OLS and TLR-OLS) consistency in ranking is assured since all observations get the same reduction or penalty for a certain characteristic.
The LGD estimates generated by OLS and RiR might also be higher than 1 or lower than 0 for certain input parameters. Since these estimates need to be truncated to the [0,1] interval a small bias may occur as seen in the MEs. In the tree models the only possible outcomes of LGD estimates are stated in the leaves. As each estimate constitutes of an average of actual LGD observations the estimates are bounded between 0 and 1. Both the LR-OLS and TLR-OLS models can potentially produce estimates smaller than 0 or larger than 1 but this is not as likely as for the OLS and RiR models since the OLS estimate only accounts for a part of the estimated value. However, in an LDP setting where only a small amount of default data is available the LR-OLS and TLR-OLS models produce rather uncertain estimates since the already small sample is divided into smaller subsets and information is removed when the regression samples are created. In this setting an OLS or RiR model gives steadier estimates of LGD as the regression uses the full sample. It might also be hard to evaluate the reasonability of the logistic regression parameters since theory might indicate whether a certain variable should influence the average LGD level positively or negatively but not whether the specific case of full or no recovery is more or less likely. This problem is potentially even more severe in a LDP setting since the need for a qualitative check of the reasonability of the parameters is more important due to the uncertainty resulting from the small sample.

<table>
<thead>
<tr>
<th>LGD (%)</th>
<th>Collateral value / Defaulted amount (%)</th>
</tr>
</thead>
<tbody>
<tr>
<td>100</td>
<td>0.0</td>
</tr>
<tr>
<td>0</td>
<td>1.43</td>
</tr>
<tr>
<td>10</td>
<td>1.43</td>
</tr>
<tr>
<td>0</td>
<td>2.310</td>
</tr>
<tr>
<td>10</td>
<td>2.311</td>
</tr>
<tr>
<td>80</td>
<td>2.312</td>
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<tr>
<td>50</td>
<td>2.313</td>
</tr>
<tr>
<td>5</td>
<td>2.313</td>
</tr>
<tr>
<td>0</td>
<td>2.620</td>
</tr>
<tr>
<td>0</td>
<td>2.622</td>
</tr>
<tr>
<td>30</td>
<td>2.67</td>
</tr>
<tr>
<td>0</td>
<td>2.68</td>
</tr>
<tr>
<td>0</td>
<td>3.01</td>
</tr>
<tr>
<td>0</td>
<td>48</td>
</tr>
<tr>
<td>10</td>
<td>72</td>
</tr>
</tbody>
</table>

Table 15. Splitting continuous variables (Illustrative figures)

The LDP setting also affects the RT model as the minimum leaf size has to be a rather large percentage of the total sample in order to include enough observations for a relatively stable estimate. In the FT model the algorithm itself handles this problem to some extent. If the difference between two groups is statistically significant then a split of the groups is performed. In this way the model finds an optimal level of splits for a given sample size. However, too small leaves could still be problematic from a practical point of view since the average LGD could change substantially if an outlier is added. In practice one would not like the estimates to change too much when new observations are incorporated.

Given that the set of explanatory variables primarily consists of dummy variables the advantage of the RT model being able to split continuous variables at the optimal place is not used to a large extent in this context. The two continuous variables used in this study, guarantee and collateral size, also both have a bimodal distribution with many observations close to 0 and 1. This might result in unintuitive and inexact splitting points as can be seen in Table 15. Even if a certain level of collateral is optimal...
to split upon from a reduction of standard deviation perspective it might not seem appropriate from a qualitative perspective. This problem is likely to be more severe in a LDP setting, given the small sample size. An alternative approach is to use a qualitative opinion to create one or more dummy variables replacing the continuous variable. The potential splitting points could then be chosen based on an expert opinion and the RT algorithm decides which ones to use.

Both tree models also give the opportunity to include expert opinions in the construction of the models. It is possible to make small deviations from the algorithms of constructing the tree and instead of using the optimal split using another split if this seems more appropriate from a qualitative perspective. In this sense the tree models could be preferred above the other models when it comes to implementing a model in practice. One should bear in mind however, that also the other models can use expert opinions when deciding e.g. which variables to include in the models and where to cap them.

When assessing the LGD levels one needs to remember the LDP setting and the small amount of data available. The Basel accord states that “the less information a bank has, the more conservative must be its assignments of exposures to borrower and facility grades or pools” (BIS, 2006, § 411).

### 6.4 Summary of Results

Some of the risk drivers believed to be suitable to include in the models have been deemed inappropriate to include for practical or mathematical reasons. The macroeconomic variables have been left out due to the difficulties of finding a variable giving the expected negative influence on the LGD estimate as well as practical concerns regarding the possibilities to produce downturn estimates by using such a variable. The seniority of the defaulted facility has, in contrast to most academic studies, been left out since it failed to prove significant in an OLS regression. This could be the result either of the small sample available or alternatively a characteristic of exposures towards financial institutions. Many academic studies use the size of the exposure as a risk driver for LGD. The exposure size has been deemed inappropriate to include partly due to the small size of the exposure in relation to the total liabilities of a global financial institution and partly due to the practical difficulties in explaining to a business unit why two small loans should differ from one large in LGD. Furthermore the utilisation rate has been considered inappropriate since it is likely to soar just before the default event and hence cannot be used to predict future LGD levels. Industry distress has not been considered since it is believed to be heavily linked to the macroeconomic environment.

From a mathematical viewpoint adjusting the observed values prior to 2005, the year when the Basel default definitions were publicized, would make the models produce better estimates. This adjustment could however be problematic to justify and might therefore not be suitable for a practical model.

All tested models produce better results than using the historical average in within and out of sample testing, both with and without adjustment of observations prior to 2005. The FT model produces the best results within sample and out of sample if the original LGD observations are used. No model produces good results in the out of time testing but the FT model is the worst and is actually worse than using the Historical average approach. On the other hand, the tree models have some advantages from a practical viewpoint. The estimates are guaranteed to lie within the \([0,1]\) interval since the estimates are averages of actual observations and it is easy to include an expert opinion when creating the models. Furthermore, the bootstrapping exercise indicates that the tree models give more stable LGD estimates for the example portfolio defined in Section 5. The TLR-OLS and particularly the LR-OLS model might be less suitable in a LDP setting since information is removed and the already small sample is divided into smaller subsets and separate regressions are performed on these
subsets. This might result in more unstable estimates. Since the LR-OLS does not produce better results than the TLR-OLS model, the TLR-OLS model might be more suitable in a LDP setting as it only divides the sample once. On a portfolio level the tested models give similar LGD estimates. However the estimates may differ substantially for specific exposures.
7 CONCLUDING REMARKS

LGD estimates should "reflect the experience and practice of the individual institution as well as the external environment in which it operates" (BIS, 2006, § 70). This implies that a financial institution cannot simply rely on general industry estimates for the parameters but need to adjust them in order to account for institution specific factors. While this is done to some extent by adjusting the LGD estimates based on the proportion of different facility types and collateral etc. the institution in question would also need to justify that the risk they are facing can be assessed by using pooled industry data, or alternatively make some adjustment.

The sample used in this study consists of pooled industry data. This means that the same default can be reported several times by different banks and the actual number of default events might be much smaller than the number of observations. Unfortunately it is hard to tell which observations belong to the same default event since the data is anonymous. Even though observations of the same default are not copies of each other they could potentially be similar which reduces the usefulness of the data set. This reduces the applicability of the out of sample testing since not all observations are independent.

Covered bonds and repos are two types of exposures common in LDPs where the lack of data is particularly severe. For these kind of exposures, a judgemental model based on expert opinions probably needs to be applied if they are to qualify for IRB treatment. This is in line with the recommendations by the British Bankers Association, the London Investment Banking Association and ISDA that no portfolios should be ruled out from IRB treatment but that it cannot be taken as a premise that all types of exposures will meet the requirements and that business experience can be used to create judgemental models if the data available is insufficient (ISDA, 2005).
8 References


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